

STAT 25100 Lecture 8

4.3 Expected Value of Discrete Random Variables

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The Expected Value of a Random Variable

Definition: Expected Value of a Discrete Random Variable

Let X be a discrete random variable with PMF $p(x)$.

The **expected value** or the **expectation** or the **mean** of X , denoted by $E[X]$, or μ_x is a **weighted average** of the possible values of X , where the weights are the probabilities of those values.

$$\mu_x = E[X] = \sum_{\text{all values of } x} x \cdot p(x)$$

provided that provided that $\sum_x |x|p(x) < \infty$.

If the sum diverges, the expectation is undefined.

Expected Value of the Geometric Distribution

Recall the Geometric PMF is

$$p(k) = (1 - p)^{x-1}p \quad \text{for } x = 1, 2, 3, \dots$$

To evaluate its expected value

$$E(X) = \sum_x x \cdot p(x) = \sum_{x=1}^{\infty} x(1 - p)^{x-1}p,$$

we'll start from the geometric series

$$\sum_{x=0}^{\infty} r^x = \frac{1}{1 - r} \quad \text{if } |r| < 1 \Rightarrow \begin{array}{l} \text{differentiate} \\ \text{both sides} \\ \text{w/ respect to } r \end{array} \Rightarrow \sum_{x=1}^{\infty} xr^{x-1} = \frac{1}{(1 - r)^2}, \quad \text{for } |r| < 1.$$

Applying the new identity with $r = 1 - p$, we get

$$E(X) = \underbrace{\sum_{x=1}^{\infty} x(1 - p)^{x-1}}_{=1/(1-(1-p))^2} p = \frac{1}{p^2} \cdot p = \frac{1}{p}.$$

Expected Value of Negative Binomial

Recall the Negative Binomial PMF is $p(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}$ $x = r, r+1, \dots$

The expected value of the Negative Binomial is

$$\begin{aligned} E(X) &= \sum_{x=r}^{\infty} x \binom{x-1}{r-1} p^r (1-p)^{x-r} \\ &= \sum_{x=r}^{\infty} r \binom{x}{r} p^r (1-p)^{x-r} \quad \left(\text{since } x \binom{x-1}{r-1} = \frac{x \cdot (x-1)!}{(r-1)!(x-r)!} = \frac{x!}{(r-1)!(x-r)!} = r \binom{x}{r} \right) \\ &= \frac{r}{p} \sum_{x=r}^{\infty} \binom{x}{r} p^{r+1} (1-p)^{x-r} \\ &= \frac{r}{p} \underbrace{\sum_{y=r+1}^{\infty} \binom{y-1}{r+1-1} p^{r+1} (1-p)^{y-(r+1)}}_{=1, \text{ since it's the sum of PMF for NB}(r+1,p)} \quad (\text{let } y = x + 1) \end{aligned}$$

r **Intuition:** (As it takes $1/p$ trials on average to get the first success,)

An Example Where $E(X)$ Is Infinite

Game: Toss a fair coin repeatedly, the longer one keeps getting heads, the more reward. Specifically, the reward is 2^n cents if one gets n consecutive Heads before the first Tail. The PMF of the reward X (in cents) is

$$P(X = 2^n) = \frac{1}{2^{n+1}}, \quad n = 0, 1, 2, \dots$$

The expected value of X is

$$E(X) = \sum_{n=0}^{\infty} 2^n P(X = 2^n) = \sum_{n=0}^{\infty} 2^n \cdot \frac{1}{2^{n+1}} = \sum_{n=0}^{\infty} \frac{1}{2} = \infty.$$

Expected Values of Functions of Random Variables

Example: Expected Values of X^2 (1)

If X is a random variable with the PMF $p_X(x)$ below, how to find the expected value of $Y = X^2$?

x	-1	0	1	2
$p_X(x) = P(X = x)$	0.3	0.4	0.2	0.1

Method 1: First we find the PMF $p_Y(y)$ of Y .

- ▶ Possible values of Y :
- ▶ $P(Y = 0) = P(X^2 = 0) = P(X = 0) = 0.4$
- ▶ $P(Y = 1) = P(X^2 = 1) =$
- ▶ $P(Y = 4) = P(X^2 = 4) = P(X = 2) = 0.1$
- ▶ PMF of Y :

y	0	1	4
$p_Y(y)$	0.4	0.5	0.1
- ▶ Expected value of $Y = X^2$ is thus

$$E(Y) = \sum_y yp_Y(y) = 0 \cdot 0.4 + 1 \cdot 0.5 + 4 \cdot 0.1 = 0.9$$

Example: Expected Values of X^2 (2)

x	-1	0	1	2
$p_X(x) = P(X = x)$	0.3	0.4	0.2	0.1

Method 2: In fact, we can find expected value of $Y = X^2$ using the PMF of X directly as

$$\begin{aligned} E(Y) &= E(X^2) = \sum_x x^2 \cdot p_X(x) \\ &= (-1)^2 \cdot 0.3 + 0^2 \cdot 0.4 + 1^2 \cdot 0.2 + 2^2 \cdot 0.1 \\ &= 0.9 \end{aligned}$$

- ▶ Same answer as Method 1
- ▶ Method 2 is easier no need to find the PMF of $Y = X^2$,

Expected Values of Functions of Random Variables

Generally, if X is a random variable with PMF $p_X(x)$, and $Y = g(X)$, how to find the expected value of Y ?

Expected Values of Functions of Random Variables

Generally, if X is a random variable with PMF $p_X(x)$, and $Y = g(X)$, how to find the expected value of Y ?

Method 1: Find the PMF $p_Y(y)$ of Y and then calculate the expected value as

$$E(Y) = \sum_y yp_Y(y)$$

Method 2: One can calculate $E(Y)$ directly using the PMF X as

$$E(Y) = E(g(X)) = \sum_x g(x)p_X(x)$$

- ▶ Method 2 is easier since one doesn't have to find the PMF of $Y = g(X)$, which is sometimes not easy

Proof of the Equivalence of the Two Methods

First note that

$$p_Y(y) = P(Y = y) = P\left(\bigcup_{x: g(x)=y} \{X = x\}\right) = \sum_{x: g(x)=y} P(X = x) = \sum_{x: g(x)=y} p_X(x)$$

Thus

$$\begin{aligned} E(Y) &= \sum_y y p_Y(y) = \sum_y y \sum_{x: g(x)=y} p_X(x) \\ &= \sum_y \sum_{x: g(x)=y} y p_X(x) \\ &= \sum_y \sum_{x: g(x)=y} g(x) p_X(x) \\ &= \sum_x g(x) p_X(x) \end{aligned}$$

Expected Value of $aX + b$

If X is a random variable, the expected value for its Linear transformation $Y = g(X) = aX + b$ is

$$E(aX + b) = a E(X) + b.$$

Proof.

$$\begin{aligned} E(aX + b) &= \sum_x (ax + b)p(x) \\ &= \sum_x (ax p(x) + bp(x)) \\ &= \sum_x axp(x) + \sum_x bp(x) \\ &= a \underbrace{\sum_x xp(x)}_{=E(X)} + b \underbrace{\sum_x p(x)}_{=1} \\ &= a E(X) + b \end{aligned}$$

$$\mathbb{E}[h_1(X) + h_2(X)] = \mathbb{E}[h_1(X)] + \mathbb{E}[h_2(X)]$$

If X is a random variable, and $g(X)$ is a function of X . Suppose $g(x)$ is the sum of two functions

$$g(x) = h_1(x) + h_2(x)$$

then

$$\mathbb{E}(g(x)) = \mathbb{E}[h_1(X) + h_2(X)] = \mathbb{E}[h_1(X)] + \mathbb{E}[h_2(X)]$$

Proof.

$$\begin{aligned}\mathbb{E}(h_1(X) + h_2(X)) &= \sum_x (h_1(x) + h_2(x))p(x) \\ &= \sum_x h_1(x)p(x) + \sum_x h_2(x)p(x) \\ &= \sum_x h_1(x)p(x) + \sum_x h_2(x)p(x) \\ &= \mathbb{E}(h_1(X)) + \mathbb{E}(h_2(X))\end{aligned}$$

Variance, Standard Deviation, and Various Moments

Variance & Standard Deviation (SD)

One measure of spread of a random variable (or its probability distribution) is the *variance*.

The **variance** of a random variable X , denoted as $\text{Var}(X)$ or σ_X^2 is defined as the **average squared distance from the expected value $\mu_X = \mathbb{E}(x)$** .

$$\begin{aligned}\text{Var}(X) &= \sigma^2 = \text{"sigma squared"} \\ &= \mathbb{E}[(X - \mu_X)^2] \\ &= \sum_x (x - \mu_X)^2 p_X(x)\end{aligned}$$

provided that the variance is $< \infty$.

Square root of the variance is the *standard deviation (SD)*.

$$\text{SD}(X) = \sigma = \sqrt{\text{Var}(X)}$$

Variance of $aX + b$

For $Y = aX + b$, we have proved that $E(Y) = E(aX + b) = a\mu + b$, where $\mu = E(X)$ and hence

$$[Y - E(Y)]^2 = [(aX + b) - E(aX + b)]^2 = [aX + b - (a\mu + b)]^2 = a^2(X - \mu)^2.$$

Variance of $aX + b$

For $Y = aX + b$, we have proved that $E(Y) = E(aX + b) = a\mu + b$, where $\mu = E(X)$ and hence

$$[Y - E(Y)]^2 = [(aX + b) - E(aX + b)]^2 = [aX + b - (a\mu + b)]^2 = a^2(X - \mu)^2.$$

Taking expected value of the above we get

$$\begin{array}{ccc} E[Y - E(Y)]^2 & = & E[a^2(X - \mu)^2] \\ \parallel & & \parallel \\ \text{Var}(Y) & & a^2 E[(X - \mu)^2] \\ \parallel & & \parallel \\ \text{Var}(aX + b) & & a^2 \text{Var}(X) \end{array}$$

This shows that

$$\text{Var}(aX + b) = a^2 \text{Var}(X).$$

Moments & Central Moments

Given a random variable X with mean μ

- ▶ its *kth moment* is defined to be $E[X^k]$, and
- ▶ its *kth central moment* is defined to be $E[(X - \mu)^k]$,

provided that $E[|X|^k] < \infty$ and $E[|X - \mu|^k] < \infty$.

Note that

- ▶ the 1st moment $E(X)$ is the mean = expected value
- ▶ the 1st central moment $E(X - \mu)$ is always 0
- ▶ the 2nd central moment $E[(X - \mu)^2]$ is the variance.

A Shortcut Formula for Calculating Variance

$$\text{Var}(X) = \text{E}[(X - \mu)^2] = \text{E}(X^2) - \mu^2$$

Proof.

$$\begin{aligned} \text{E}[(X - \mu)^2] &= \sum_x (x - \mu)^2 p(x) \\ &= \sum_x (x^2 - 2\mu x + \mu^2) p(x) \\ &= \underbrace{\sum_x x^2 p(x)}_{=\text{E}(X^2)} - 2\mu \underbrace{\sum_x x p(x)}_{=\mu} + \mu^2 \underbrace{\sum_x p(x)}_{=1} \\ &= \text{E}(X^2) - 2\mu^2 + \mu^2 = \text{E}(X^2) - \mu^2 \end{aligned}$$

Example of Variance

Find the variance for a random variable X with the PMF below.

x	-1	0	1	2
$p_X(x)$	0.3	0.4	0.2	0.1

Sol #1: Use the definition of variance.

$$\mu = E(X) = \sum_x xp_X(x) = (-1) \cdot 0.3 + 0 \cdot 0.4 + 1 \cdot 0.2 + 2 \cdot 0.1 = 0.1$$

$$\begin{aligned}\text{Var}(X) &= E(X - \mu)^2 = \sum_x (x - 0.1)^2 p_X(x) \\ &= (-1 - 0.1)^2 \cdot 0.3 + (0 - 0.1)^2 \cdot 0.4 + (1 - 0.1)^2 \cdot 0.2 + (2 - 0.1)^2 \cdot 0.1 \\ &= 1.21 \cdot 0.3 + 0.01 \cdot 0.4 + 0.81 \cdot 0.2 + 3.61 \cdot 0.1 \\ &= 0.89\end{aligned}$$

x	-1	0	1	2
$p_X(x)$	0.3	0.4	0.2	0.1

Sol #2: Use the shortcut formula.

$$E(X) = \sum_x xp_X(x) = (-1) \cdot 0.3 + 0 \cdot 0.4 + 1 \cdot 0.2 + 2 \cdot 0.1 = 0.1$$

$$E(X^2) = \sum_x x^2 \cdot p_X(x)$$

$$= (-1)^2 \cdot 0.3 + 0^2 \cdot 0.4 + 1^2 \cdot 0.2 + 2^2 \cdot 0.1 = 0.9$$

$$\text{Var}(X) = E(X^2) - \mu^2 = 0.9 - 0.1^2 = 0.89$$

▶ same answer & easier to calculate!

Factorial Moments

Given a random variable X with mean μ , its *kth moment* is defined to be

$$\mathbb{E}[X(X - 1) \cdots (X - k + 1)], \quad k = 1, 2, 3, \dots$$

- ▶ first factorial moment = $\mathbb{E}(X)$ = first moment
- ▶ For Binomial, Hypergeometric, and Poisson, Factorial Moments are easier to calculate than Moments

Factorial Moments of Poisson

Recall the Poisson PMF is $p(x) = e^{-\lambda} \frac{\lambda^x}{x!}$, $x = 0, 1, 2, \dots$

$$\begin{aligned} & E[X(X-1)\cdots(X-k+1)] \\ &= \sum_{x=0}^{\infty} x(x-1)\cdots(x-k+1)e^{-\lambda} \frac{\lambda^x}{x!} \\ &= \sum_{x=k}^{\infty} x(x-1)\cdots(x-k+1)e^{-\lambda} \frac{\lambda^x}{x!} \quad \text{first } k \text{ terms are 0} \\ &= \sum_{x=k}^{\infty} e^{-\lambda} \frac{\lambda^x}{(x-k)!} \\ &= \lambda^k e^{-\lambda} \sum_{x=k}^{\infty} \frac{\lambda^{x-k}}{(x-k)!} \\ &= \lambda^k e^{-\lambda} \underbrace{\sum_{y=0}^{\infty} \frac{\lambda^y}{y!}}_{=e^\lambda} \quad \text{let } y = x - k \\ &= \lambda^k \end{aligned}$$

Mean & Variance of Poisson

For $X \sim \text{Poisson}(\lambda)$, the k th factorial moments are

$$\mathbb{E}[X(X-1)\cdots(X-k+1)] = \lambda^k, \quad k = 1, 2, 3, \dots$$

- ▶ $k = 1$: $\mathbb{E}(X) = \lambda$
- ▶ $k = 2$: $\mathbb{E}(X(X-1)) = \lambda^2$
- ▶ To find the 2nd moment $\mathbb{E}(X^2)$, observe that $X^2 = X(X-1) + X$. Thus

$$\mathbb{E}(X^2) = \mathbb{E}(X(X-1)) + \mathbb{E}(X) = \lambda^2 + \lambda$$

- ▶ $\text{Var}(X) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2 = \lambda^2 + \lambda - \lambda^2 = \lambda$.

Factorial Moments of Binomial (2)

Recall the Binomial PMF is $p(x) = \binom{n}{x} p^x (1-p)^{n-x}$, $0 \leq x \leq n$.

$$\begin{aligned} & \mathbb{E}[X(X-1)\cdots(X-k+1)] \\ &= \sum_{x=0}^n x(x-1)\cdots(x-k+1) \binom{n}{x} p^x (1-p)^{n-x} \\ &= \sum_{x=k}^n x(x-1)\cdots(x-k+1) \binom{n}{x} p^x (1-p)^{n-x} \quad \text{first } k \text{ terms are } 0 \end{aligned}$$

Key step:

$$\begin{aligned} x(x-1)\cdots(x-k+1) \binom{n}{x} &= x(x-1)\cdots(x-k+1) \frac{n!}{x!(n-x)!} \\ &= \frac{n!}{(x-k)!(n-x)!} \\ &= n(n-1)\cdots(n-k+1) \binom{n-k}{x-k}. \end{aligned}$$

Factorial Moments of Binomial (2)

$$\begin{aligned}
 & E[X(X-1)\cdots(X-k+1)] \\
 & \sum_{x=k}^n x(x-1)\cdots(x-k+1) \binom{n}{x} p^x (1-p)^{n-x} \\
 & = \sum_{x=k}^n n(n-1)\cdots(n-k+1) \binom{n-k}{x-k} p^x (1-p)^{n-x} \quad (\text{key step in previous page}) \\
 & = n(n-1)\cdots(n-k+1)p^k \sum_{x=k}^n \binom{n-k}{x-k} p^{x-k} (1-p)^{n-x} \\
 & = n(n-1)\cdots(n-k+1)p^k \underbrace{\sum_{y=0}^{n-k} \binom{n-k}{y} p^y (1-p)^{n-k-y}}_{=(p+1-p)^{n-k}=1} \quad \text{let } y = x - k \\
 & = n(n-1)\cdots(n-k+1)p^k \left(\begin{array}{l} \text{Binomial} \\ \text{expansion} \end{array} (a+b)^N = \sum_{i=0}^N \binom{N}{i} a^i b^{N-i} \text{ with } \begin{array}{l} a = p \\ b = 1-p \end{array} \right)
 \end{aligned}$$

Mean & Variance of Binomial

For $X \sim \text{Bin}(n, p)$, the k th factorial moments are

$$\mathbb{E}[X(X-1)\cdots(X-k+1)] = n(n-1)\cdots(n-k+1)p^k, \quad k = 1, 2, 3, \dots$$

- ▶ $k = 1$: $\mathbb{E}(X) = np$
- ▶ $k = 2$: $\mathbb{E}(X(X-1)) = n(n-1)p^2$
- ▶ To find the 2nd moment $\mathbb{E}(X^2)$, observe that $X^2 = X(X-1) + X$. Thus

$$\mathbb{E}(X^2) = \mathbb{E}(X(X-1)) + \mathbb{E}(X) = n(n-1)p^2 + np$$

- ▶
$$\text{Var}(X) = \mathbb{E}(X^2) - (\mathbb{E}(X))^2 = n(n-1)p^2 + np - (np)^2 = np - np^2 = np(1-p).$$

Factorial Moments of Hypergeometric

Hypergeometric PMF:

$$p(x) = \binom{R}{x} \binom{B}{d-x} / \binom{R+B}{d}, \quad 0 \leq x \leq R, \quad 0 \leq d-x \leq B.$$

Factorial Moment:

$$E[X(X-1)\cdots(X-k+1)] = \sum_{x=0}^R x(x-1)\cdots(x-k+1) \binom{R}{x} \binom{B}{d-x} / \binom{R+B}{d}.$$

We first calculate its numerator on the next page

$$\begin{aligned}
& \sum_{x=0}^R x(x-1)\cdots(x-k+1) \binom{R}{x} \binom{B}{d-x} \\
&= \sum_{x=k}^R x(x-1)\cdots(x-k+1) \binom{R}{x} \binom{B}{d-x} \\
&= \sum_{x=k}^R R(R-1)\cdots(R-k+1) \binom{R-k}{x-k} \binom{B}{d-x} \quad (\text{same key step as Binomial}) \\
&= R(R-1)\cdots(R-k+1) \sum_{i=0}^{R-k} \binom{R-k}{i} \binom{B}{d-k-i} \quad (\text{let } i = x - k) \\
&= R(R-1)\cdots(R-k+1) \binom{R+B-k}{d-k} \quad (\text{Vandermonde's identity})
\end{aligned}$$

Putting back the denominator $\binom{R+B}{d}$, we get the factorial moment

$$\mathbb{E}[X(X-1)\cdots(X-k+1)] = R(R-1)\cdots(R-k+1) \binom{R+B-k}{d-k} / \binom{R+B}{d}$$

Mean & Variance of Hypergeometric

For $X \sim \text{Hypergeometric}(R, B, d)$, the k th factorial moments, $k = 1, 2, 3, \dots$, are

$$\begin{aligned} \mathbb{E}[X(X-1)\cdots(X-k+1)] &= R(R-1)\cdots(R-k+1) \binom{R+B-k}{d-k} / \binom{R+B}{d} \\ &= \frac{R(R-1)\cdots(R-k+1) \times d(d-1)\cdots(d-k+1)}{(R+B)(R+B-1)\cdots(R+B-k+1)} \end{aligned}$$

► $k = 1$: $\mathbb{E}(X) = \frac{dR}{R+B}$

► $k = 2$: $\mathbb{E}(X(X-1)) = \frac{R(R-1)d(d-1)}{(R+B)(R+B-1)}$

► To find the 2nd moment $\mathbb{E}(X^2)$, observe that $X^2 = X(X-1) + X$. Thus

$$\mathbb{E}(X^2) = \mathbb{E}(X(X-1)) + \mathbb{E}(X) = \frac{R(R-1)d(d-1)}{(R+B)(R+B-1)} + \frac{dR}{R+B}$$

$$\begin{aligned} \text{Var}(X) &= \mathbb{E}(X^2) - [\mathbb{E}(X)]^2 \\ &= \frac{R(R-1)d(d-1)}{(R+B)(R+B-1)} + \frac{Rd}{R+B} - \frac{R^2d^2}{(R+B)^2} = d \frac{RB(R+B-d)}{(R+B)^2(R+B-1)}. \end{aligned}$$

Variance for Negative Binomial (1)

Negative Binomial PMF: $p(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r} \quad x = r, r+1, \dots$

For Negative Binomial, $E(X(X+1))$ is easier to calculate than $E(X^2)$ or $E(X(X-1))$.

$$E(X(X+1)) = \sum_{x=r}^{\infty} x(x+1) \binom{x-1}{r-1} p^r (1-p)^{x-r}$$

Key step:

$$\begin{aligned} x(x+1) \binom{x-1}{r-1} &= \frac{(x+1)x \cdot (x-1)!}{(x-r)!(r-1)!} = \frac{(x+1)!}{(x-r)!(r-1)!} = \frac{r(r+1) \cdot (x+1)!}{(x-r)!(r+1)!} \\ &= r(r+1) \binom{x+1}{r+1}. \end{aligned}$$

Variance for Negative Binomial (2)

$$\begin{aligned}E(X(X+1)) &= \sum_{x=r}^{\infty} x(x+1) \binom{x-1}{r+1} p^r (1-p)^{x-r} \\&= \sum_{x=r}^{\infty} r(r+1) \binom{x+1}{r+1} p^r (1-p)^{x-r} \quad (\text{key step on previous page}) \\&= \frac{r(r+1)}{p^2} \sum_{x=r}^{\infty} \binom{x+1}{r+1} p^{r+2} (1-p)^{x-r} \\&= \frac{r(r+1)}{p^2} \underbrace{\sum_{y=r+2}^{\infty} \binom{y-1}{r+2-1} p^{r+2} (1-p)^{y-(r+2)}}_{=1, \text{ since it's the sum of PMF for NB}(r+2,p)} \quad (\text{let } y = x + 2) \\&= \frac{r(r+1)}{p^2}.\end{aligned}$$

Variance for Negative Binomial (3)

The second moment:

$$E(X^2) = E(X(X + 1)) - E(X) = \frac{r(r + 1)}{p^2} - \frac{r}{p}.$$

By the shortcut formula, the variance is

$$\text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{r(r + 1)}{p^2} - \frac{r}{p} - \frac{r^2}{p^2} = \frac{r(1 - p)}{p^2}$$

Summary: Common Discrete Distributions — Definitions

Distribution	Genesis
Bernoulli (p)	1 for Success, 0 for Failure, for the outcome of one Bernoulli trial
Binomial(n, p)	number of successes obtained in n Bernoulli trials, that each trial has a success probability p
Geometric(p)	number of Bernoulli trials needed to get the first success, that each trial has a success probability p
Negative Binomial(r, p)	number of Bernoulli trials needed to get r successes, that each trial has a success probability p
Hypergeometric (R, B, d)	number of red marbles obtained in d draws without replacement from a box with R red and B blue marbles.
Poisson(λ)	limit of Binomial(n, p) as $n \rightarrow \infty$, $p \rightarrow 0$, and $np \rightarrow \lambda$

Summary: Common Discrete Distributions — PMF, Mean, Variance

Name and range	PMF at x	Mean	Variance
Bernoulli(p) on $\{0, 1\}$	$\begin{cases} 1 - p & \text{if } x = 0 \\ p & \text{if } x = 1 \end{cases}$	p	$p(1 - p)$
Binomial(n, p) on $\{0, 1, \dots, n\}$	$\binom{n}{x} p^x (1 - p)^{n-x}$	np	$np(1 - p)$
Geometric(p) on $\{1, 2, 3 \dots\}$	$(1 - p)^{x-1} p$	$\frac{1}{p}$	$\frac{1 - p}{p^2}$
Negative Binomial(r, p) on $\{r, r + 1, r + 2, \dots\}$	$\binom{x-1}{r-1} p^r (1 - p)^{x-r}$	$\frac{r}{p}$	$\frac{r(1 - p)}{p^2}$
Hypergeometric(R, B, d) on $\{0, 1, 2, \dots, \min(d, R)\}$	$\frac{\binom{R}{x} \binom{B}{d-x}}{\binom{R+B}{d}}$	$d \frac{R}{R+B}$	$d \frac{RB(R+B-d)}{(R+B)^2(R+B-1)}$
Poisson(λ) on $\{0, 1, 2, \dots\}$	$e^{-\lambda} \frac{\lambda^x}{x!}$	λ	λ

Summary: Relations Between Discrete Distributions

► Bernoulli & Binomial

► $\text{Binomial}(n = 1, p) = \text{Bernoulli}(p)$

► If X_1, \dots, X_n are i.i.d. $\text{Bernoulli}(p)$, then $\sum_{i=1}^n X_i \sim \text{Binomial}(n, p)$

► If $X \sim \text{Binomial}(m, p)$ and $Y \sim \text{Binomial}(n, p)$ are independent, then $X + Y \sim \text{Binomial}(m + n, p)$

► Geometric & Negative Binomial

► $\text{Negative Binomial}(r = 1, p) = \text{Geometric}(p)$

► If X_1, \dots, X_n are i.i.d. $\text{Geometric}(p)$, then $\sum_{i=1}^n X_i \sim \text{Negative Binomial}(r = n, p)$

► If $X \sim \text{NegBin}(m, p)$ and $Y \sim \text{NegBin}(n, p)$ are independent, then $X + Y \sim \text{NegBin}(m + n, p)$

► If $R \gg d$ and $B \gg d$,

$$\text{Hypergeometric}(R, B, d) \approx \text{Binomial}\left(n = d, p = \frac{R}{R + B}\right)$$

► If $n \rightarrow \infty$, $p \rightarrow 0$, and $np \rightarrow \lambda$, $\text{Binomial}(n, p) \approx \text{Poisson}(\lambda = np)$

► If $X_i \sim \text{Poisson}(\lambda_i)$ are independent, then $\sum_{i=1}^n X_i \sim \text{Poisson}(\sum_{i=1}^n \lambda_i)$