STAT 234 Lecture 13 Central Limit Theorem Section 6.1-6.2

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- the average of random sample $\{X_1, \ldots, X_n\}$, $\overline{X} = \frac{1}{n}(X_1 + \cdots + X_n)$ is called the *sample mean*
- Observe that the sample mean X is also a random variable, which has a probability distribution, called the sampling distribution of the (sample) mean.

In Lectured 11, we showed if X_1, \ldots, X_n are **i.i.d.** random variables with *mean* μ and *variance* σ^2 , then

$$E(\overline{X}) = \mu$$
, $Var(\overline{X}) = \frac{\sigma^2}{n}$

from which we can prove the Weak Law of Large Numbers:

as
$$n \to \infty$$
, $\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i \to \mu$.

Intuitively, this is clear from the mean and the variance of \overline{X} ; the "center" of the distribution \overline{X} is μ , and the "spread" around it becomes smaller and smaller as *n* grows.

Note that the sample mean \overline{X} itself is a random variable, and hence it has a probability distribution, called the *sampling distribution of the (sample) mean*.

The sampling distribution of \overline{X} depends on the **population** distribution. Here are some examples.

- If $X_1, \ldots, X_n \sim N(\mu, \sigma^2)$, then $\overline{X} \sim N(\mu, \sigma^2/n)$.
- If \overline{X} is the average of *n* Bernoulli random variables $X_1, \ldots, X_n \sim \text{Bernoulli}(p)$, then $n\overline{X} \sim Bin(n, p)$, i.e.,

$$P\left(\overline{X} = \frac{k}{n}\right) = \binom{n}{k} p^k (1-p)^{n-k}, \quad 0 \le k \le n.$$

and so on.

Let $X_1, X_2, ...$ be **i.i.d.** random variables with *mean* μ and *variance* σ^2 . CLT asserts that, when *n* is large,

• the distribution of the sample mean $\overline{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$ is approximately

$$N\left(\mu_{\overline{x}}=\mu, \ \sigma_{\overline{x}}^2=\frac{\sigma^2}{n}\right).$$

• the distribution of the *total* $T = \sum_{i=1}^{n} X_i$ is approximately

$$N\left(\mu_T = n\mu, \ \sigma_T^2 = n\sigma^2\right).$$

Recall the <u>card game</u> in Lecture 4, draw ONE card from a well-shuffled deck of cards and get a reward based on the card drawn as follows.

Event	reward X	p(x)
Heart (not ace)	\$1	12/52
Ace	\$5	4/52
King of spades	\$10	1/52
All else	\$0	35/52
Total		1

- The card drawn is placed back to the deck before he draws the card for the next game.
- Let *X_i* be the reward he get in the *i*th game, then *X_i*'s are i.i.d. and his total reward from the 300 games is

$$X_1 + X_2 + \dots + X_{300}$$
 6

Recall the pmf for the reward X_i from one game is

The expected reward from one game and the variance are

$$\mu = \mathcal{E}(X) = 0 \cdot \frac{35}{52} + 1 \cdot \frac{12}{52} + 5 \cdot \frac{4}{52} + 10 \cdot \frac{1}{52} = \frac{21}{26}$$
$$\mathcal{E}(X^2) = 0^2 \cdot \frac{35}{52} + 1^2 \cdot \frac{12}{52} + 5^2 \cdot \frac{4}{52} + 10^2 \cdot \frac{1}{52} = \frac{53}{13}$$
$$\sigma^2 = \operatorname{Var}(X) = \mathcal{E}(X^2) - \mu^2 = \frac{53}{13} - \left(\frac{21}{26}\right)^2 = \frac{2315}{26^2}$$

So if a gambler played the game 300 times, his expected value, variance of his total reward is

$$E(X_1 + \dots + X_{300}) = 300\mu = 300 \times \frac{21}{26} \approx 243.308$$
$$Var(X_1 + \dots + X_{300}) = 300\sigma^2 = 300 \times \frac{2315}{26^2}$$
$$SD(X_1 + \dots + X_{300}) = \sqrt{300 \times \frac{2315}{26^2}} = 32.052$$

The gambler is expected to get \$243.308 from the 300 games, with a standard deviation \$32.052.

Example 1: Card Game

What is the probability that the gambler can earn \$250 or more from the 300 games?

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Solution: By CLT, as n = 300 is large, the distribution of the total rewards $T = \sum_{i=1}^{300} X_i$ is approx. normal w/

 $\mu_T = n\mu = 300\mu = 243.308, \quad \sigma_T = \sqrt{300}\sigma = 32.052.$

Thus

$$P(\text{total reward} > \$250) = P\left(Z > \frac{250 - 243.308}{32.052}\right)$$
$$\approx P(Z > 0.21) \approx 1 - 0.5832 \approx 0.417$$

1- pnorm(250, m = 243.308, s = 32.052) [1] 0.4173 Suppose a company ships packages that vary in weight:

- Packages have mean 15 lb and standard deviation 10 lb.
- Packages weights are independent from each other

Q: What is the probability that the average weight of 100 packages exceeds 17 lb?

Let W_i be the weight of the *i*th package and the total weights of 100 packages is

$$\overline{W} = \frac{1}{100} \sum_{i=1}^{100} W_i,$$

where W_i 's are i.i.d. with mean $\mu_W = 15$ and SD $\sigma_W = 10$.

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By CLT, \overline{W} is approx. $N(\mu_{\overline{w}} = 15, \sigma_{\overline{w}}^2 = 1^2)$,

$$P(\overline{W} > 17) = P\left(\frac{\overline{W} - \mu_{\overline{W}}}{\sigma_{\overline{W}}} > \frac{17 - \mu_{\overline{W}}}{\sigma_{\overline{W}}}\right)$$
$$= P\left(Z > \frac{17 - 15}{1}\right) \approx 1 - \Phi(2) \approx 0.023$$

1- pnorm(2) [1] 0.02275

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, for $x > 0$, $\mu = 1$, $\sigma^2 = 1$



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blue curve: the normal approximation



Normal Approximation to Binomial Distribution

Normal approximation to the Binomial distributions is a special case of CLT:

$$X = \sum_{i=1}^{n} X_i \sim Bin(n, p),$$

where $X_1, X_2, ..., X_n$ are *n* independent Bernoulli random variables with success probability *p*.

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By CLT, for large $n, Y \sim Bin(n, p)$ is approximately distributed as

$$N(\mu_Y = np, \ \sigma_Y^2 = np(1-p)).$$

Normal Approximation to Bin(n, p = 0.5)

When $X_1, \ldots, X_n \sim \text{Bernoulli}(p = 0.5)$, the sampling distribution of \overline{X} is







If the population distribution is skewed, so is the sampling distribution of the sample mean, though the skewness diminishes as the number of draws goes up. With a perfectly balanced roulette wheel, red numbers should turn up 18 in 38 of the time. To test its wheel, one casino records the results of 3800 plays. Let X be the number of reds the casino got.

Q1: If the roulette wheel is perfectly balanced, what is the chance that $X \ge 1890$?

Q2 If the casino gets 1890 reds, do you think the roulette wheel should be calibrated?



Example 3: Roulette Calibration

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Thus

$$E(X) = np = 3800(18/38) = 1800$$

$$SD(X) = \sqrt{np(1-p)} = \sqrt{3800(18/38)(20/38)} \approx 30.78$$
By CLT, X is approximately $N(\mu = 1800, \sigma^2 = (30.78)^2)$. Thus,
$$P(X \ge 1800) \approx p\left(\frac{X - 1800}{2} \ge \frac{1890 - 1800}{2}\right) \approx p(Z \ge 2.02) \approx 0.001$$

$$P(X \ge 1890) \approx P\left(\frac{X - 1800}{30.78} \ge \frac{1890 - 1800}{30.78}\right) \approx P(Z \ge 2.92) \approx 0.00173$$

1-pnorm(1890, m = 1800, s = sqrt(3800*(18/38)*(20/38))) [1] 0.001728 As $X \sim \text{Bin}(n = 3800, p = 18/38)$, the exact probability of $X \ge 1890$ is

$$P(X \ge 1890) = \sum_{k=1890}^{3800} \binom{3800}{k} \left(\frac{18}{38}\right)^k \left(\frac{20}{38}\right)^{3800-k} \approx 0.00183$$

found using R as follows.

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We can see normal approx. to Binomial gives fairly good approx to the exact Binomial probability. As $X \sim \text{Bin}(n = 3800, p = 18/38)$, the exact probability of $X \ge 1890$ is

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Q2 If the casino gets 1890 reds, do you think the roulette wheel should be calibrated? Yes. $X \ge 1890$ is very unlikely to happen.

- If the population is normal, then any *n* will do.
- If the population distribution is symmetric, then *n* should be at least 30 or so.
- The more skew or irregular the population, the larger *n* has to be
- For the Binomial distribution, a rule of thumb is that *n* should be such that

$$np \ge 10$$
 and $n(1-p) \ge 10$.