# STAT 234 Lecture 5 Binomial Distributions Section 3.5 

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## Coverage

Lecture 5 covers Section 3.5 of MMSA
Please skip

- Moment General Function for Binomial on p. 135 in Section 3.5
- Section 3.4 Moment Generating Functions
- Section 3.6 Hypergeometric and Negative Binomial Distributions
- Section 3.7 Poisson Distributions


## The Binomial Distribution

Five draws are made at random with replacement from a box containing one red ball and 9 green balls.


What is the probability that the first two draws are Red and the next 3 are Green?

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$$
\begin{aligned}
\mathrm{P}(R R G G G) & =\mathrm{P}(R) \cdot \mathrm{P}(\mathrm{R}) \cdot \mathrm{P}(\mathrm{G}) \cdot \mathrm{P}(\mathrm{G}) \cdot \mathrm{P}(\mathrm{G}) \\
& =0.1 \times 0.1 \times 0.9 \times 0.9 \times 0.9
\end{aligned}
$$

As the draws are made with replacement, the outcomes of the 5 draws are independent.
The multiplication rule for independent events can be applied.

What is the probability of getting exactly two Reds in 5 draws? Is it also equal to

$$
0.1 \times 0.1 \times 0.9 \times 0.9 \times 0.9 ?
$$

There are 10 possible orderings of the 2 Reds and the 3 Greens.

| Possible Orders | Probability |
| :---: | :---: |
| RRG G G | $0.1 \times 0.1 \times 0.9 \times 0.9 \times 0.9=(0.1)^{2}(0.9)^{3}$ |
| RGRGG | $0.1 \times 0.9 \times 0.1 \times 0.9 \times 0.9=(0.1)^{2}(0.9)^{3}$ |
| RGGRG | $0.1 \times 0.9 \times 0.9 \times 0.1 \times 0.9=(0.1)^{2}(0.9)^{3}$ |
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| G G GRR | $0.9 \times 0.9 \times 0.9 \times 0.1 \times 0.1=(0.1)^{2}(0.9)^{3}$ |

P (exactly 2 Reds in 5 draws) is the sum of the probabilities of the 10 cases above because the 10 cases are disjoint. So
$\mathrm{P}($ exactly 2 Reds in 5 draws $)=10 \times(0.1)^{2}(0.9)^{3}$.

What is P (getting exactly $k$ Reds in $n$ draws)?

What is P (getting exactly $k$ Reds in $n$ draws)?
Consider all possible ways to order the $k$ Reds and the $n-k$ Greens.


Note

- the events for different orderings are disjoint, and
- each occurs with identical probability $(0.1)^{k}(0.9)^{n-k}$.

By the Addition Rule, P (exactly $k$ Reds in $n$ draws) equals
(\# of ways to order $k$ Reds and $n-k$ Greens) $\times(0.1)^{k}(0.9)^{n-k}$

## Factorial

The notation $n$ !, read $n$ factorial, is defined as

$$
n!=1 \times 2 \times 3 \times \ldots \times(n-1) \times n
$$

e.g.,

$$
\begin{array}{ll}
1!=1, & 3!=1 \times 2 \times 3=6 \\
2!=1 \times 2=2, & 4!=1 \times 2 \times 3 \times 4=24
\end{array}
$$

By convention,

$$
0!=1
$$

## Binomial Coefficients

The number of ways to order $k$ Reds and $n-k$ Greens equals

$$
\binom{n}{k}=\frac{n!}{k!(n-k)!}
$$

- $\binom{n}{k}$ is read as " $n$ choose $k$ ", also denoted as ${ }_{n} C_{k}$, or $C_{k}^{n}$. e.g.,

$$
\begin{aligned}
& \binom{5}{2}=\frac{5!}{2!\times(5-2)!}=\frac{5 \times 4 \times 3 \times 2 \times 1}{(2 \times 1)(3 \times 2 \times 1)}=\frac{5 \times 4}{2 \times 1}=10, \\
& \binom{n}{n}=\frac{n!}{n!\times 0!}=\frac{n!}{n!\times 1}=1
\end{aligned}
$$

You can also use R for these calculations:
choose $(5,2)$
[1] 10

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$$
\begin{aligned}
& \cdot\left({ }_{(0)}^{(0)}=\right. \\
& \cdot\left({ }_{(0)}^{(0)}=\right.
\end{aligned}
$$

$$
-\left(i_{i}^{(i)}=\right.
$$

$$
\cdot\left(n_{n-1}^{n}\right)=
$$

- $\binom{n}{0}=\frac{n!}{0!\times n!}=1$
$\Rightarrow$ there is only 1 way to order 0 Reds and $n$ Greens
-(i) $=$
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- $\left(\begin{array}{l}n-1\end{array}\right)=$
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- $\binom{n}{n}=\frac{n!}{n!\times 0!}=1$
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- $\binom{n}{1}=\frac{n!}{1!\times(n-1)!}=n$
$\Rightarrow$ there are $n$ ways to order 1 Red and $n-1$ Greens
- $\binom{n}{n-1}=$
- $\binom{n}{0}=\frac{n!}{0!\times n!}=1$
$\Rightarrow$ there is only 1 way to order 0 Reds and $n$ Greens
- $\binom{n}{n}=\frac{n!}{n!\times 0!}=1$
$\Rightarrow$ there is only 1 way to order $n$ Reds and 0 Green
- $\binom{n}{1}=\frac{n!}{1!\times(n-1)!}=n$
$\Rightarrow$ there are $n$ ways to order 1 Red and $n-1$ Greens
- $\binom{n}{n-1}=\frac{n!}{(n-1)!\times 1!}=n$
$\Rightarrow$ there are $n$ ways to order $n-1$ Reds and 1 Green


## Summary of All the Calculations So Far

When $n$ draws are made at random with replacement from a box contains one red ball and 9 green ones,

the probability to get exactly $k$ Reds (and $n-k$ Greens) equals

$$
\binom{n}{k}(0.1)^{k}(0.9)^{n-k}
$$

Such calculations can be generalized to other similar problems and the general formula is called the Binomial Formula.

## Bernoulli Trials

A random trial having only 2 possible outcomes (Success, Failure) is called a Bernoulli trial, e.g.,

- whether a coin lands heads or tails when tossing a coin
- whether one gets a six or not a six when rolling a die
- whether a drug works on a patient or not
- whether a electronic device is defected
- whether a subject answers Yes or No to a survey question


## Binomial Formula

Suppose $n$ independent Bernoulli trials are to be performed, each of which results in

- a success with probability $p$ and
- a failure with probability $1-p$.

The probability of getting $k$ successes and $n-k$ failures in $n$ Bernoulli trials is given by
(\# of ways to order the $k$ successes and $n-k$ failures) $\times p^{k}(1-p)^{n-k}$
$=\binom{n}{k} p^{k}(1-p)^{n-k}$

## Binomial Distribution

Suppose $n$ independent Bernoulli trials are to be performed, each of which results in

- a success with probability $p$ and
- a failure with probability $1-p$.

If we define

$$
X=\text { the number of successes that occur in the } n \text { trials, }
$$

then $X$ is said to have a binomial distribution with parameters ( $n, p$ ), denoted as

$$
X \sim \operatorname{Bin}(n, p)
$$

with the probability mass function (pmf)

$$
\mathrm{P}(X=k)=\binom{n}{k} p^{k}(1-p)^{n-k}, \quad k=0,1, \ldots, n
$$

## Does the Binomial Probabilities Add Up to 1?

Recall the Binomial expansion in math:

$$
(a+b)^{n}=\sum_{k=0}^{n}\binom{n}{k} a^{k} b^{n-k}
$$

which is valid for all real numbers $a$ and $b$.
Applying the Binomial expansion with $a=p$ and $b=1-p$, we get

$$
\sum_{k=0}^{n} \mathrm{P}(X=k)=\sum_{k=0}^{n}\binom{n}{k} p^{k}(1-p)^{n-k}=(p+1-p)^{n}=1^{n}=1
$$

which means that the pmf

$$
\mathrm{P}(X=k)=\binom{n}{k} p^{k}(1-p)^{n-k}, \quad k=0,1, \ldots, n
$$

is a valid pmf.

## Conditions Required to be Binomial

Conditions required to apply the binomial formula:

1. each trial outcome must be classified as a success or a failure
2. the probability of success, $p$, must be the same for each trial
3. the number of trials, $n$, must be fixed
4. the trials must be independent

## Binomial or Not? - 10 Rolls of a Die

Rolling a die 10 times, what is the probability of getting exactly 3 aces?

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- prob. of success $p=1 / 6$
- number of trials $n=10$


## Binomial or Not? - 10 Rolls of a Die

Rolling a die 10 times, what is the probability of getting exactly 3 aces?

- a trial: whether one gets an ace when rolling a die once
- prob. of success $p=1 / 6$
- number of trials $n=10$
- the trials (rolls) are independent

So, it's okay to use the Binomial formula.

$$
\begin{aligned}
\mathrm{P}(3 \text { aces in } 10 \text { rolls }) & =\frac{10!}{3!7!}\left(\frac{1}{6}\right)^{3}\left(1-\frac{1}{6}\right)^{7} \\
& =\frac{10 \times 9 \times 8 \times(7!)}{(3 \times 2 \times 1)(7!)}\left(\frac{1}{6}\right)^{3}\left(\frac{5}{6}\right)^{7} \\
& =120\left(\frac{1}{6}\right)^{3}\left(\frac{5}{6}\right)^{7} \approx 0.155
\end{aligned}
$$

## Binomial or Not? — Rolling a Die Until 3rd Ace

Rolling a die continuously, is the probability of getting the 3rd aces in the 10th roll equals to

$$
\frac{10!}{3!7!}\left(\frac{1}{6}\right)^{3}\left(1-\frac{1}{6}\right)^{7} ?
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$$

No. The number of trials (sample size) is not determined in advance.

## Binomial or Not? — Restaurants

Suppose an inspector randomly selects 5 restaurantsfrom the 20 restaurants in a town, of which 10 currently have health code violation(s) and the other 10 have no violations. Let $X$ be the number of selected restaurants with violations. Is $X$ binomial?

- a trial: whether a randomly selected restaurant has violation(s)

$$
S=\text { violation }, \quad F=\text { no violation }
$$

- number of trials: $n=5$


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- number of trials: $n=5$
- prob. of success?


## Binomial or Not? — Restaurants (2)

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- $P(S$ on $2 n d$ trial | $F$ on 1 st trial $)=\frac{10}{19}$
- $P(S$ on 5 th trial $\mid F F F F)=\frac{10}{16}=0.625$


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- $P(S$ on 2 nd trial | F on 1 st trial $)=\frac{10}{19}$
- $P(S$ on 5 th trial | FFFF $)=\frac{10}{16}=0.625$
- $P(S$ on 5 th trial $\mid S S S S)=\frac{6}{16}=0.375$


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- $P(S$ on 5 th trial | FFFF $)=\frac{10}{16}=0.625$
- $P(S$ on 5 th trial | SSSS $)=\frac{6}{16}=0.375$
- Trials are NOT independent since the selection are made without replacement.


## Binomial or Not — UC Undergrads

50 UC undergrads are randomly selected and each of them is asked whether he/she has a driver's licence. Suppose 4000 of the 5000 UC undergrads have driver's licence. Let $X$ be the number who reply yes. Is $X$ binomial?

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- a trial: a randomly selected student reply yes $=\mathrm{S}$ or no $=\mathrm{F}$
- number of trials $n=50$


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- a trial: a randomly selected student reply yes $=$ S or no $=F$
- number of trials $n=50$
- Strictly speaking, NOT binomial, since selection are made without replacement - trials are dependent. However,...
- $P(S$ on first trial $)=\frac{4000}{5000}=0.8$
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- $\mathrm{P}(\mathrm{S}$ on 50 th trial | all S up to 49 th $)=\frac{3951}{4951} \approx 0.7980$
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- Since the sample size 50 is only $1 \%$ of the population size (5000), the 50 draws has little effect on the makeup of the population. $P(S)$ stays close to 0.8 regardless of the outcome of prior draws. Trials are nearly independent
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- Since the sample size 50 is only $1 \%$ of the population size (5000), the 50 draws has little effect on the makeup of the population. $P(S)$ stays close to 0.8 regardless of the outcome of prior draws. Trials are nearly independent
- So $X$ is approx. binomial, $\operatorname{Bin}(n=50, p=0.8)$.


## Rule: Approx. Binomial or Not When Sampling w/o Replace-

 mentConsider sampling without replacement from a dichotomous population of size $N$. If the sample size (number of trials) $n$ is at most $5 \%$ of the population size, the experiment can be analyzed as though it were exactly a binomial experiment.

