## STAT 226 Lecture 25

Section 8.1 Comparing Dependent Proportions

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## Example (Two Case Control Studies on Hodgkin's Disease)

Study 1 by Diana et al. (1971) :

| Tonsillectomy |  |  |  |
| :--- | :---: | :---: | :---: |
|  | Yes | No | Total |
| Cases | 67 | 34 | 101 |
| Controls | 43 | 64 | 109 |

$$
\begin{aligned}
\widehat{\theta}=\frac{67 \times 64}{34 \times 43} & \approx 2.93 \\
\text { Pearson's } X^{2} & =13.23 \\
d f & =1 \\
P \text {-value } & =0.00028
\end{aligned}
$$

Study 2 by Johnson \& Johnson (1972):

| Tonsillectomy |  |  |  |
| :--- | :---: | :---: | :---: |
|  | Yes | No | Total |
| Cases | 41 | 44 | 85 |
| Controls | 33 | 52 | 85 |

$$
\widehat{\theta}=\frac{41 \times 52}{44 \times 33} \approx 1.47
$$

Pearson's $X^{2}=1.17$

$$
d f=1
$$

$$
P \text {-value }=0.279
$$

Why did the 2 studies reach inconsistent conclusions?

- In study 1, controls were matched as a group according to age, sex, race, residence, date of hospitalization, and no cancer
- In study 2, controls were patient's same sex sibling closest in age. To reflect dependence, data in Study 2 should be displayed as

| d as |  | Sibling had tonsillectomy? |  | Total |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Yes | No |  |
| Hodgkin's patient | Yes | 26 | 15 | 41 |
| had tonsillectomy? | No | 7 | 37 | 44 |
|  | Total | 33 | 52 | 85 |

This table shows high correlation of tonsillectomy status within a sibling pair. Most pairs of siblings had the same tonsillectomy status.

- concordant pairs tell us nothing about the relationship between the disease and the risk factor. Sufficient to examine discordant pairs only

Population probabilities:

|  | Sibling had <br> tonsillectomy? |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
|  | Yes |  |  |  |
| No | Total |  |  |  |
|  | Yes | $\pi_{11}$ | $\pi_{12}$ | $\pi_{1+}$ |
|  | No | $\pi_{21}$ | $\pi_{22}$ | $\pi_{2+}$ |
|  | Total | $\pi_{+1}$ | $\pi_{+2}$ | 1 |

## Discussion

- What was the goal of the study?
- Which two proportions were we interested in comparing?

Population probabilities:

|  |  | Sibling had tonsillectomy? |  | Total |
| :---: | :---: | :---: | :---: | :---: |
|  |  | Yes | No |  |
| Hodgkin's patient | Yes | $\pi_{11}$ | $\pi_{12}$ | $\pi_{1+}$ |
| had tonsillectomy? | No | $\pi_{21}$ | $\pi_{22}$ | $\pi_{2+}$ |
|  | Total | $\pi_{+1}$ | $\pi_{+2}$ | 1 |

Discussion

- What was the goal of the study?
- Which two proportions were we interested in comparing?

Compare dependent samples by making inference about $\pi_{1+}-\pi_{+1}$. There is marginal homogeneity if $\pi_{1+}=\pi_{+1}$.

## McNemar's Test

Under $\mathrm{H}_{0}$ : marginal homogeneity,

$$
\pi_{1+}=\pi_{+1} \Longleftrightarrow \pi_{12}=\pi_{21} \Longleftrightarrow \frac{\pi_{12}}{\pi_{12}+\pi_{21}}=\frac{1}{2}
$$

Under $\mathrm{H}_{0}$, each of $n^{*}=n_{12}+n_{21}$ observations has probability $1 / 2$ of contributing to $n_{12}$ and $1 / 2$ of contributing to $n_{21}$ :

$$
n_{12} \sim \operatorname{Binomial}\left(n^{*}, \frac{1}{2}\right), \text { mean }=\frac{n^{*}}{2}, \text { std dev }=\sqrt{n^{*}\left(\frac{1}{2}\right)\left(\frac{1}{2}\right)}
$$

By normal approx. to binomial, for large $n^{*}$,

$$
z=\frac{n_{12}-n^{*} / 2}{\sqrt{n^{*}\left(\frac{1}{2}\right)\left(\frac{1}{2}\right)}}=\frac{n_{12}-n_{21}}{\sqrt{n_{12}+n_{21}}} \sim N(0,1)
$$

or equivalently

$$
z^{2}=\frac{\left(n_{12}-n_{21}\right)^{2}}{n_{12}+n_{21}} \sim \chi_{1}^{2}
$$

## Example (Two Case Control Studies on Hodgkin's Disease)

## Hodgkin's patient had tonsillectomy?

Sibling had

|  | tonsillectomy? |  |  |  |
| :--- | :---: | ---: | :---: | :---: |
|  |  | Yes | No | Total |
| Hodgkin's patient | Yes | 26 | 15 | 41 |
| had tonsillectomy? | No | 7 | 37 | 44 |
|  | Total | 33 | 52 | 85 |

For testing $\mathrm{H}_{0}: \pi_{1+}=\pi_{+1}$ v.s. $\mathrm{H}_{a}: \pi_{1+} \neq \pi_{+1}$

$$
\begin{aligned}
z=\frac{n_{12}-n_{21}}{\sqrt{n_{12}+n_{21}}} & =\frac{15-7}{\sqrt{15+7}}=1.706, \quad\left(z^{2}=2.909, d f=1\right) \\
P \text {-value } & =0.088 .
\end{aligned}
$$

Though still insignificant, Study 2 is more consistent w/ Study 1.

## McNemar's Test in $\mathbf{R}$

hodgkin1972 $=$ matrix $(c(26,7,15,37)$, nrow=2)
hodgkin1972
[,1] [,2]
[1,] $26 \quad 15$
[2,] $7 \quad 37$
mcnemar.test(hodgkin1972, correct = FALSE)

McNemar's Chi-squared test
data: hodgkin1972
McNemar's chi-squared $=2.9091, \mathrm{df}=1, \mathrm{p}$-value $=0.08808$

## Exact McNemar's Test

When $n^{*}=n_{12}+n_{21}$ are small, it's more reliable to use exact binomial tests since under $\mathrm{H}_{0}: \pi_{12}=\pi_{21}$, the exact distribution of
$n_{12}$ is

$$
n_{12} \sim \operatorname{Binomial}\left(n^{*}, \frac{1}{2}\right)
$$

For the Hodgkin's Disease study, the exact two-sided $P$-value is binom.test(15, 22, p=0.5)

## Exact binomial test

```
data: }15\mathrm{ and 22
```

number of successes $=15$, number of trials $=22, \mathrm{p}$-value $=0.1338$
alternative hypothesis: true probability of success is not equal to 0.5
95 percent confidence interval:
0.45127560 .8613535
sample estimates:
probability of success
0.6818182

## Exact McNemar's Test

$P$-values given by McNemar's test with continuity correction are closer to the $P$-values given by exact McNemar's test.
mcnemar.test(hodgkin1972, correct $=$ TRUE)
McNemar's Chi-squared test with continuity correction
data: hodgkin1972
McNemar's chi-squared $=2.2273, \mathrm{df}=1, \mathrm{p}$-value $=0.1356$

## CI for $\pi_{1+}-\pi_{+1}$

Estimate $\pi_{1+}-\pi_{+1}$ by diff. of sample proportions, $\widehat{\pi}_{1+}-\widehat{\pi}_{+1}$.

$$
\begin{aligned}
& \widehat{\pi}_{1+}-\widehat{\pi}_{+1}=\frac{n_{1+}}{n}-\frac{n_{+1}}{n}=\frac{n_{12}-n_{21}}{n} \\
& \mathrm{SE}=\frac{1}{n} \sqrt{n_{12}+n_{21}-\frac{\left(n_{12}-n_{21}\right)^{2}}{n}}
\end{aligned}
$$

## Example (Hodgkin's Disease)

| $n_{11}$ | $n_{12}$ |  |
| :---: | :---: | :--- |
| $n_{21}$ | $n_{22}$ |  |
|  |  | $n$ |$=$| 26 | 15 |  |
| :---: | :---: | :---: |
| 7 | 37 |  |

$$
\begin{gathered}
\widehat{\pi}_{1+}-\widehat{\pi}_{+1}=\frac{15-7}{85} \approx 0.094 \\
\mathrm{SE}=\frac{1}{85} \sqrt{15+7-\frac{(15-7)^{2}}{85}} \approx 0.054
\end{gathered}
$$

$95 \%$ CI: $0.094 \pm(1.96)(0.054)=(-0.0118,0.1998)$

## How is the SE of the Matched-Pairs CI Derived?

$$
\begin{gathered}
\left(n_{11}, n_{12}, n_{21}, n_{22}\right) \sim \operatorname{Multinomial}\left(n,\left(\pi_{11}, \pi_{12}, \pi_{21}, \pi_{22}\right)\right) \\
\Longrightarrow\left\{\begin{array}{l}
\operatorname{Var}\left(n_{i j}\right)=n \pi_{i j}\left(1-\pi_{i j}\right) \\
\operatorname{Cov}\left(n_{i j}, n_{i^{\prime} j^{\prime}}\right)=-n \pi_{i j} \pi_{i^{\prime} j^{\prime}} \quad \text { if } i \neq i^{\prime} \text { or } j \neq j^{\prime}
\end{array}\right. \\
\begin{aligned}
\operatorname{Var}\left(\widehat{\pi}_{1+}-\widehat{\pi}_{+1}\right) & =\operatorname{Var}\left(\frac{n_{12}-n_{21}}{n}\right)=\frac{\operatorname{Var}\left(n_{12}-n_{21}\right)}{n^{2}} \\
& =\frac{\operatorname{Var}\left(n_{12}\right)+\operatorname{Var}\left(n_{21}\right)-2 \operatorname{Cov}\left(n_{12}, n_{21}\right)}{n^{2}} \\
& =\frac{n \pi_{12}\left(1-\pi_{12}\right)+n \pi_{21}\left(1-\pi_{21}\right)+2 n \pi_{12} \pi_{21}}{n^{2}} \\
& =\frac{\pi_{12}+\pi_{21}-\left(\pi_{12}^{2}-2 \pi_{12} \pi_{21}+\pi_{21}^{2}\right)}{n} \\
& =\frac{\pi_{12}+\pi_{21}-\left(\pi_{12}-\pi_{21}\right)^{2}}{n} \quad(\mathrm{ctd} \text { next slide) }
\end{aligned}
\end{gathered}
$$

## How is the SE of the Matched-Pairs CI Derived? (Cont'd)

$$
\begin{aligned}
\operatorname{Var}\left(\widehat{\pi}_{1+}-\widehat{\pi}_{+1}\right) & =\frac{\pi_{12}+\pi_{21}-\left(\pi_{12}-\pi_{21}\right)^{2}}{n} \\
\widehat{\operatorname{Var}}\left(\widehat{\pi}_{1+}-\widehat{\pi}_{+1}\right) & =\frac{\widehat{\pi}_{12}+\widehat{\pi}_{21}-\left(\widehat{\pi}_{12}-\widehat{\pi}_{21}\right)^{2}}{n} \\
& =\frac{\frac{n_{12}}{n}+\frac{n_{21}}{n}-\left(\frac{n_{12}}{n}-\frac{n_{21}}{n}\right)^{2}}{n} \\
& =\frac{\frac{n_{12}}{n}+\frac{n_{21}}{n}-\frac{\left(n_{12}-n_{21}\right)^{2}}{n^{2}}}{n} \times \frac{n}{n} \\
& =\frac{n_{12}+n_{21}-\frac{\left(n_{12}-n_{21}\right)^{2}}{n}}{n^{2}}
\end{aligned}
$$

So

$$
\mathrm{SE}=\sqrt{\operatorname{Var}\left(\widehat{\pi}_{1+}-\widehat{\pi}_{+1}\right)}=\frac{1}{n} \sqrt{n_{12}+n_{21}-\frac{\left(n_{12}-n_{21}\right)^{2}}{n}}
$$

## Two-Sample Designs v.s Matched-Pair Designs

$$
\begin{aligned}
& \operatorname{Var}\left(\widehat{\pi}_{1+}-\widehat{\pi}_{+1}\right)=\operatorname{Var}\left(\widehat{\pi}_{1+}\right)+\operatorname{Var}\left(\widehat{\pi}_{+1}\right)-2 \operatorname{Cov}\left(\widehat{\pi}_{1+}, \widehat{\pi}_{+1}\right) \\
& \text { where } \operatorname{Var}\left(\widehat{\pi}_{1+}\right)=\frac{\pi_{1+}\left(1-\pi_{1+}\right)}{n}, \operatorname{Var}\left(\widehat{\pi}_{+1}\right)=\frac{\pi_{+1}\left(1-\pi_{+1}\right)}{n}, \\
& \operatorname{Cov}\left(\widehat{\pi}_{1+}, \widehat{\pi}_{+1}\right)=\operatorname{Cov}\left(\frac{n_{1+}}{n}, \frac{n_{+1}}{n}\right)=\operatorname{Cov}\left(\frac{n_{11}+n_{12}}{n}, \frac{n_{11}+n_{21}}{n}\right) \\
&= \frac{1}{n^{2}} \operatorname{Cov}\left(n_{11}+n_{12}, n_{11}+n_{21}\right) \\
&= \frac{1}{n^{2}}\left[\operatorname{Var}\left(n_{11}\right)+\operatorname{Cov}\left(n_{11}, n_{21}\right)+\operatorname{Cov}\left(n_{12}, n_{11}\right)+\operatorname{Cov}\left(n_{12}, n_{21}\right)\right] \\
&= \frac{1}{n^{2}}\left[n \pi_{11}\left(1-\pi_{11}\right)-n \pi_{11} \pi_{21}-n \pi_{12} \pi_{11}-n \pi_{12} \pi_{21}\right] \\
&= \frac{1}{n}[\pi_{11} \underbrace{\left(1-\pi_{11}-\pi_{12}-\pi_{21}\right)}_{\pi_{22}}-\pi_{12} \pi_{21}] \\
&= \frac{\pi_{11} \pi_{22}-\pi_{12} \pi_{21}}{n}
\end{aligned}
$$

## Two-Sample Designs v.s Matched-Pair Designs (Cont’d)

Often matched-pairs exhibit positive association (odds-ratio greater than 1), i.e., $\pi_{11} \pi_{22}>\pi_{12} \pi_{21}$, so covariance term is negative.

$$
\begin{aligned}
& \operatorname{Var}\left(\widehat{\pi}_{1+}-\widehat{\pi}_{+1}\right) \text { for matched-pairs design } \\
= & \frac{1}{n}[\pi_{1+}\left(1-\pi_{1+}\right)+\pi_{+1}\left(1-\pi_{+1}\right)-2(\underbrace{\pi_{11} \pi_{22}-\pi_{12} \pi_{21}}_{\text {usually }>0})] \\
\leq & \frac{1}{n}\left[\pi_{1+}\left(1-\pi_{1+}\right)+\pi_{+1}\left(1-\pi_{+1}\right)\right] \\
= & \operatorname{Var}\left(\widehat{\pi}_{1+}-\widehat{\pi}_{+1}\right) \text { for two indep. samples of size } n \text { each }
\end{aligned}
$$

If matched-pairs exhibit positive association, matched-pairs designs are more efficient than 2-sample designs.

## Example (Opinions Relating to Environment)

In 2000 General Social Survey, subjects were asked whether, to help the environment, they would be willing to (1) pay higher taxes or (2) accept a cut in living standards.

|  | Yes | No | Total | \% Yes |
| :--- | :---: | :---: | :---: | :---: |
| Pay Higher Taxes? | 334 | 810 | 1144 | $29.2 \%$ |
| Cut Living Standards? | 359 | 785 | 1144 | $31.4 \%$ |

- Q: Which option were there more people willing to accept?
- The two sample percentages are dependent because the same subjects were asked both questions. There were 1144 subjects only, not $1144+1144$.

To reflect dependence, data should be displayed as

|  | Willing to Cut |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
|  | Living Standards? |  |  |  |
|  |  | Yes | No | Total |
| Willing to Pay | Yes | 227 | 132 | 359 |
| Higher Taxes? | No | 107 | 678 | 785 |
|  | Total | 334 | 810 | 1144 |

## Example (Opinions Relating to Environment)

|  | Willing to Cut |  |  |  |
| :--- | :---: | :---: | :---: | :---: |
|  |  | Living Standards? |  |  |
|  |  | Yes | No | Total |
| Willing to Pay | Yes | 227 | 132 | 359 |
| Higher Taxes? | No | 107 | 678 | 785 |
|  | Total | 334 | 810 | 1144 |

Estimates for $\pi_{1+}=$ proportion willing to pay higher taxes, and $\pi_{+1}=$ proportion willing to cut living standards, are respectively

$$
\widehat{\pi}_{1+}=\frac{359}{1144} \approx 0.314, \quad \text { and } \widehat{\pi}_{+1}=\frac{334}{1144} \approx 0.292
$$

The $95 \%$ confidence interval for $\pi_{1+}-\pi_{+1}$ is NOT
$(0.314-0.292) \pm 1.96 \sqrt{\frac{0.314(1-0.314)}{1144}+\frac{0.292(1-0.292)}{1144}} \approx 0.022 \pm 0.038$
The correct Cl is

$$
(0.314-0.292) \pm 1.96 \frac{1}{1144} \sqrt{132+107-\frac{(132-107)^{2}}{1144}} \approx 0.022 \pm 0.026
$$

## Example (Opinions Relating to Environment)

|  | Table I |  |  | Table II |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Pay <br> Higher <br> Taxes? | Cut Living Standards? |  | Total |
|  | Yes | No | Total |  | Yes | No |  |
| Pay Higher Taxes? | 334 | 810 | 1144 | Yes | 227 | 132 | 359 |
| Cut Living Standards? | 359 | 785 | 1144 | No | 107 | 678 | 785 |
|  |  |  |  | Total | 334 | 810 | 1144 |

To test whether there were more people willing to pay higher taxes or more people willing to cut living standards, we should ...
a. perform Pearson's $X^{2}$ test on Table I
b. perform Pearson's $X^{2}$ test on Table II
c. perform McNemar's test on Table I
d. perform McNemar's test on Table II

Which one is correct?

## Example (Opinions Relating to Environment)

|  | Table I |  |  | Table II |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Pay Higher Taxes? | Cut Living Standards? |  | Total |
|  | Yes | No | Total |  | Yes | No |  |
| Pay Higher Taxes? | 334 | 810 | 1144 | Yes | 227 | 132 | 359 |
| Cut Living Standards? | 359 | 785 | 1144 | No | 107 | 678 | 785 |
|  |  |  |  | Total | 334 | 810 | 1144 |

To test whether those willing to pay higher taxes were more or less willing to cut living standards than those not willing to pay higher taxes, we should ...
a. perform Pearson's $X^{2}$ test on Table I
b. perform Pearson's $X^{2}$ test on Table II
c. perform McNemar's test on Table I
d. perform McNemar's test on Table II

Which one is correct?

