

February 18, 2013

7.4 Renewal Reward Processes  
7.5.1 Alternating Renewal Processes

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Proof of Proposition 7.3(a)

We give the proof for (a) only. To prove this, write

$$\frac{R(t)}{t} = \frac{\sum_{n=1}^{N(t)} R_n}{t} = \frac{\sum_{n=1}^{N(t)} R_n}{N(t)} \times \frac{N(t)}{t}$$

By the strong law of large numbers we obtain

$$\frac{\sum_{n=1}^{N(t)} R_n}{N(t)} \rightarrow \mathbb{E}[R_1] \text{ as } t \rightarrow \infty$$

and by Proposition 7.1

$$\frac{N(t)}{t} \rightarrow \frac{1}{\mathbb{E}[X_1]} \text{ as } t \rightarrow \infty$$

The result thus follows.

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Example 7.12 (A Car Buying Model) Solution

- ▶ An event occurs whenever Mr. Brown buys a new car
- ▶ Interarrival times:  $X_i = \min(Y_i, T)$
- ▶ Cost incurred in the  $i$ th cycle:  $R_i = C_1 + C_2 \mathbf{1}_{\{Y_i \leq T\}}$
- ▶ Are  $(X_i, R_i)$ ,  $i = 1, 2, \dots$  i.i.d?

$$\mathbb{E}[X_i] = \int_0^\infty \min(y, T)h(y)dy = \int_0^T yh(y)dy + T(1 - H(T))$$

- ▶  $\mathbb{E}[R_i] = C_1 + C_2P(Y_i \leq T) = C_1 + C_2H(T)$
- ▶ long-run average cost

$$= \frac{C_1 + C_2H(T)}{\int_0^T yh(y)dy + T(1 - H(T))}$$

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7.4 Renewal Reward Processes

Let  $\{N(t), t \geq 0\}$  be a renewal process with i.i.d. interarrival times  $X_i$ ,  $i = 1, 2, \dots$ . Let  $R_i$ ,  $i = 1, 2, \dots$  be i.i.d random variables.  $R_i$  may depend on the  $i$ th interarrival time  $X_i$ , but  $(X_i, R_i)$  are i.i.d. random variable pairs. The compound process

$$R(t) = \sum_{i=1}^{N(t)} R_i$$

is called a *renewal reward process*.  $R_i$  may be considered as reward earned during the  $i$ th cycle, and  $R(t)$  represents the total reward earned up to time  $t$ .

**Proposition 7.3** If  $\mathbb{E}[R_1] < \infty$  and  $\mathbb{E}[X_1] < \infty$ , then

(a) with probability 1,  $\lim_{t \rightarrow \infty} \frac{R(t)}{t} = \frac{\mathbb{E}[R_1]}{\mathbb{E}[X_1]}$

(b)  $\lim_{t \rightarrow \infty} \frac{\mathbb{E}[R(t)]}{t} = \frac{\mathbb{E}[R_1]}{\mathbb{E}[X_1]}$

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Example 7.12 (A Car Buying Model)

- ▶ Mr. Brown buy a new car whenever his old one breaks down or reaches the age of  $T$  years
- ▶ Let  $Y_i$  be the lifetime of his  $i$ th car. Suppose  $Y_i$ 's are i.i.d with CDF

$$H(y) = P(Y \leq y), \text{ and density } h(y) = H'(y).$$

- ▶ Cost to buy a new car =  $C_1$ ;
- ▶ If the car breaks down, an additional cost of  $C_2$  is incurred.
- ▶ What's Mr. Brown's long run average cost?

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Example 7.16 & 7.17

Let  $\{N(t), t \geq 0\}$  be a renewal process with i.i.d. interarrival times  $X_i$ ,  $i = 1, 2, \dots$ . Consider the current age of the item in use at time  $t$

$$A(t) = t - S_{N(t)}.$$

What is the long-run average of age

$$\lim_{t \rightarrow \infty} \frac{\int_0^t A(s)ds}{t} ?$$

Also consider the residual life of the item in use at time  $t$

$$Y(t) = S_{N(t)+1} - t.$$

What is the long-run average of residual life

$$\lim_{t \rightarrow \infty} \frac{\int_0^t Y(s)ds}{t} ?$$

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### Solution to Example 7.16

Let's try to turn  $\int_0^t A(s)ds$  into a renewal reward process:

Note  $\int_0^{S_{N(t)}} A(s)ds \leq \int_0^t A(s)ds < \int_0^{S_{N(t)+1}} A(s)ds$ , and

$$\begin{aligned} \int_0^{S_{N(t)}} A(s)ds &= \sum_{i=1}^{N(t)} \int_{S_{i-1}}^{S_i} A(s)ds = \sum_{i=1}^{N(t)} \int_{S_{i-1}}^{S_i} s - S_{i-1} ds \\ &= \sum_{i=1}^{N(t)} \int_0^{X_i} u du \quad (\text{let } u = s - S_{i-1}) \\ &= \sum_{i=1}^{N(t)} \frac{X_i^2}{2} = R(t), \end{aligned}$$

where  $R(t) = \sum_{i=1}^{N(t)} R_i$  is a renewal reward process with  $R_i = X_i^2/2$ . Similarly, one can show that

$$\int_0^{S_{N(t)+1}} A(s)ds = \sum_{i=1}^{N(t)+1} R_i = R(t) + R_{N(t)+1}.$$

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### Solution to Example 7.16 (Cont'd)

Since

$$R(t) \leq \int_0^t A(s)ds < R(t) + R_{N(t)+1},$$

and

$$\frac{R_{N(t)+1}}{t} = \frac{X_{N(t)+1}^2}{2t} \rightarrow 0 \quad \text{as } t \rightarrow \infty,$$

by Proposition 7.3, the long-run average age of the item in use is

$$\lim_{t \rightarrow \infty} \frac{\int_0^t A(s)ds}{t} = \lim_{t \rightarrow \infty} \frac{R(t)}{t} = \frac{\mathbb{E}[R_1]}{\mathbb{E}[X_1]} = \frac{\mathbb{E}[X_1^2]}{2\mathbb{E}[X_1]}$$

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### Solution to Example 7.17

Similarly, from that

$$\int_0^{S_{N(t)}} Y(s)ds \leq \int_0^t Y(s)ds < \int_0^{S_{N(t)+1}} Y(s)ds,$$

one can show that

$$\begin{aligned} \int_0^{S_{N(t)}} Y(s)ds &= \sum_{i=1}^{N(t)} \int_{S_{i-1}}^{S_i} (S_i - s)ds = \sum_{i=1}^{N(t)} \int_0^{X_i} u du \quad (\text{let } u = S_i - s) \\ &= \sum_{i=1}^{N(t)} \frac{X_i^2}{2} = R(t) \end{aligned}$$

and that  $\int_0^{S_{N(t)+1}} Y(s)ds = \sum_{i=1}^{N(t)+1} \frac{X_i^2}{2} = R(t) + R_{N(t)+1}$ . By the same argument, the long-run average of residual life of the item in use is

$$\lim_{t \rightarrow \infty} \frac{\int_0^t Y(s)ds}{t} = \lim_{t \rightarrow \infty} \frac{R(t)}{t} = \frac{\mathbb{E}[R_1]}{\mathbb{E}[X_1]} = \frac{\mathbb{E}[X_1^2]}{2\mathbb{E}[X_1]}$$

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### 7.5.1. Alternating Renewal Processes

Considers a system that can be in one of two states: **ON** or **OFF**. Initially it is ON, and remains ON for a time  $Z_1$ ; it then goes OFF and remains OFF for a time  $Y_1$ . It then goes ON for a time  $Z_2$ ; then OFF for a time  $Y_2$ ; then on, and so on. Suppose  $(Z_k, Y_k)$  are i.i.d random vectors, though  $Z_k$  and  $Y_k$  might depend on each other, and we assume  $Y_k, Z_k$  are non-negative with finite means. Then a renewal process  $\{N(t), t \geq 0\}$  with interarrival times

$$X_k = Z_k + Y_k, \quad k \geq 1$$

is called an *alternating renewal process*.

Let

$$U(t) = \begin{cases} 1 & \text{if the system is ON at time } t \\ 0 & \text{otherwise} \end{cases}$$

Q1: What is the long-run proportion of time that the system is on?

$$\lim_{t \rightarrow \infty} \frac{\int_0^t U(s)ds}{t} ?$$

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### 7.5.1. Alternating Renewal Processes (Cont'd)

The alternating renewal process can be regarded as a reward process with reward  $R_i = Z_i$ ,

$$R(t) = \sum_{i=1}^{N(t)} Z_i$$

Then

$$R(t) \leq \int_0^t U(s)ds < R(t) + Z_{N(t)+1}$$

By Proposition 7.3, with probability 1,

$$\lim_{t \rightarrow \infty} \frac{R(t)}{t} = \frac{\mathbb{E}[Z_1]}{\mathbb{E}[X_1]} = \frac{\mathbb{E}[Z_1]}{\mathbb{E}[Z_1] + \mathbb{E}[Y_1]}$$

and hence

$$\lim_{t \rightarrow \infty} \frac{\int_0^t U(s)ds}{t} = \lim_{t \rightarrow \infty} \frac{R(t)}{t} = \frac{\mathbb{E}[Z_1]}{\mathbb{E}[Z_1] + \mathbb{E}[Y_1]} = \frac{\mathbb{E}[\text{ON}]}{\mathbb{E}[\text{ON}] + \mathbb{E}[\text{OFF}]}$$

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### Definition: A Lattice Distribution

A random variable  $X$  is said to have a *lattice distribution* if there is an  $h > 0$  for which

$$\sum_{k=-\infty}^{\infty} P(X = kh) = 1$$

in which the largest  $h$  is called the *span* of  $X$

Example 1. Many discrete distributions, like Poisson, Binomial, are lattice distributions.

Example 2. Continuous distributions are non-lattice. Mixtures of discrete and continuous distributions are also non-lattice.

Remark: If  $X_i$ 's are i.i.d with a common lattice distribution, then

$$S_n = X_1 + \dots + X_n$$

also has a lattice distribution for all  $n$

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**Theorem:** If the distribution of the interarrival times is non-lattice, then

$$\lim_{t \rightarrow \infty} P(\text{the system is on at time } t) = \lim_{t \rightarrow \infty} P(U(t) = 1) = \frac{\mathbb{E}[Z_1]}{\mathbb{E}[Z_1] + \mathbb{E}[Y_1]}$$

**Exercise 7.39**

- ▶ Two independent machines, each functions for an exponential time with rate  $\lambda$
- ▶ A single repairmen. All repair times are independent with distribution function  $G$
- ▶ If the repairmen is free when a machine fails, he will begin repairing that machine immediately; Otherwise, then that machine must wait until the other machine has been repaired.
- ▶ Once repaired, a machine is as good as new.
- ▶ What proportion of time is the repairmen idle?

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**Example 7.23 & 7.24**

Let  $\{N(t), t \geq 0\}$  be a renewal process with i.i.d. interarrival times  $X_i, i = 1, 2, \dots$ , where  $\mu = \mathbb{E}[X_i]$  and  $F(x) = P(X_i \leq x)$ . Consider the current age of the item in use at time  $t$

$$A(t) = t - S_{N(t)},$$

and the residual life of the item in use at time  $t$

$$Y(t) = S_{N(t)+1} - t.$$

**Proposition** The long-run proportion of time that  $A(t) \leq x$  is the same as the long-run proportion of time that  $Y(t) \leq x$ , and is

$$F_e(x) = \frac{1}{\mu} \int_0^x (1 - F(u)) du.$$

Furthermore, if  $F$  is non-lattice,

$$\lim_{t \rightarrow \infty} P(A(t) \leq x) = \lim_{t \rightarrow \infty} P(Y(t) \leq x) = F_e(x)$$

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**Example 7.24 (Con'd)**

For  $Y(t)$ ,

- ▶ let's say the system is OFF at time  $t$  if  $Y(t) \leq x$
- ▶ length of OFF time  $Z_i = \min(X_i, x)$

$$\mathbb{E}[Z_i] = \mathbb{E}[\min(X_i, x)] = \int_0^x (1 - F(u)) du$$

- ▶ length of a cycle =  $X_i, \mathbb{E}[\text{ON}] + \mathbb{E}[\text{OFF}] = \mathbb{E}[X_i] = \mu$
- ▶ long-run proportion of time that  $Y(t) \leq x$

$$\frac{\mathbb{E}[\text{OFF}]}{\mathbb{E}[\text{ON}] + \mathbb{E}[\text{OFF}]} = \frac{1}{\mu} \int_0^x (1 - F(u)) du$$

**Remark 1:** The ON time in Example 7.23 is not the same as the ON time in Example 7.24

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**Exercise 7.39 Solutions**

- ▶ The system is ON when the repairmen is idling, OFF when busy
- ▶ length of ON time:  $Z \sim \text{Exp}(2\lambda), \mathbb{E}[Z] = 1/(2\lambda)$
- ▶ length of OFF time  $Y, \mathbb{E}[Y] = ?$
- ▶  $T =$  the time it takes to repair the first failing machine.  $T \sim G$
- ▶  $U =$  the time the working machine can function after the first machine failed. By the memoryless property,  $U \sim \text{Exp}(\lambda)$
- ▶ Note that

$$Y = T + Y' \mathbf{1}_{\{T > U\}}$$

where  $Y'$  is the time the repairmen remains busy after the first failing machine is fixed. Note that  $Y'$  is independent of  $T$  and  $U$ , and has the same distribution as  $Y$ . Thus

$$\mathbb{E}[Y] = \mathbb{E}[T] + \mathbb{E}[Y]P(T > U) \Rightarrow \mathbb{E}[Y] = \frac{\mathbb{E}[T]}{P(T < U)}$$

- ▶ long-run proportion of ON time =

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**Example 7.23 (Con'd)**

For  $A(t)$ ,

- ▶ let's say the system is ON at time  $t$  if  $A(t) \leq x$
- ▶ length of ON time  $Y_i = \min(X_i, x)$

$$\begin{aligned} \mathbb{E}[Y_i] &= \mathbb{E}[\min(X_i, x)] = \int_0^\infty P(\min(X_i, x) > u) du \\ &= \int_0^x (1 - F(u)) du \end{aligned}$$

- ▶ length of a cycle =  $X_i, \mathbb{E}[\text{ON}] + \mathbb{E}[\text{OFF}] = \mathbb{E}[X_i] = \mu$
- ▶ long-run proportion of time that  $A(t) \leq x$

$$\frac{\mathbb{E}[\text{ON}]}{\mathbb{E}[\text{ON}] + \mathbb{E}[\text{OFF}]} = \frac{1}{\mu} \int_0^x (1 - F(u)) du$$

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**About  $F_e$**

The density and  $k$ th moment of the distribution  $F_e$  is

$$f_e(x) = \frac{1}{\mu}(1 - F(x)), \text{ and } \int_0^\infty x^k f_e(x) dx = \frac{\mathbb{E}[X^{k+1}]}{(k+1)\mathbb{E}[X]}$$

Recall that

$$\frac{m(t)}{t} = \frac{1}{\mu} - \frac{1}{t} + \frac{\mathbb{E}[Y(t)]}{t\mu}.$$

If  $F$  is non-lattice, since the limiting distribution of  $Y(t)$  is  $F_e$ , we have

$$\lim_{t \rightarrow \infty} \mathbb{E}[Y(t)] = \frac{\mu^2 + \sigma^2}{2\mu}$$

Thus

$$\begin{aligned} m(t) &= \frac{t}{\mu} - 1 + \frac{\mu^2 + \sigma^2}{2\mu^2} + o(t) \\ &= \frac{t}{\mu} + \frac{\sigma^2 - \mu^2}{2\mu^2} + o(t) \end{aligned}$$

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