

P AND P^*

$(B_t, S_t^{(t)}, \dots, S_t^{(p)})$: securities

P : actual probability

P^* : risk neutral probability

Relationship: mutual absolute continuity $P \sim P^*$

For example:

$$P : dS_t = \mu_t S_t dt + \sigma_t S_t dW_t$$

$$P^* : dS_t = \mu_t^* S_t dt + \sigma_t^* S_t dW_t^*$$

Money market bond numeraire:

$$\mu_t^* = r_t$$

$$Q : \sigma_t^* = ?? \quad A : \sigma_t^{*2} = \sigma_t^2. \text{ Why?}$$

CONTINUOUS CASE: σ_t AND σ_t^*

$$P : dS_t = \mu_t S_t dt + \sigma_t S_t dW_t$$

$$P^* : dS_t = \mu_t^* S_t dt + \sigma_t^* S_t dW_t^*$$

$$P : d \log S_t = \dots dt + \sigma_t dW_t$$

hence:
$$\underbrace{(d \log S_t)^2}_{d[\log S, \log S]_t} = \sigma_t^2 (dW_t)^2 = \sigma_t^2 dt$$

$$P^* : \text{ same argument : } d[\log S, \log S]_t = \sigma_t^{*2} dt$$

Process same under P, P^* :

$$\sigma_t^{*2} dt = d[\log S, \log S]_t = \sigma_t^2 dt$$

$$\text{hence: } \sigma_t^{*2} = \sigma_t^2$$

In particular: if σ is constant:

$$\sigma^* = \sigma$$

If numeraire = Money Market Bond:

$$dS_t = r_t S_t dt + \sigma_t S_t dW_t^*$$

$$d\tilde{S}_t = \sigma_t \tilde{S}_t dW_t^*$$

CONTINUOUS CASE: THE MARKET PRICE OF RISK

$$dS_t = \mu_t S_t dt + \sigma_t S_t dW_t$$

$$=$$

$$dS_t = \mu_t^* S_t dt + \sigma_t S_t dW_t^*$$

$$\mu_t S_t dt + \sigma_t S_t dW_t = \mu_t^* S_t dt + \sigma_t S_t dW_t^*$$

$$\text{hence: } \mu_t dt + \sigma_t dW_t = \mu_t^* dt + \sigma_t dW_t^*$$

$$\text{hence: } \underbrace{\frac{\mu_t - \mu_t^*}{\sigma_t}}_{\lambda_t} dt + dW_t = dW_t^*$$

Change from P to P^* :

Market price of risk

$$\underbrace{dW_t^*}_{P^* - BM} = \underbrace{dW_t}_{P - BM} + \lambda_t dt$$

OTHER NUMERAIRE

$$\left. \begin{aligned} d \log S_t &= \dots dt + \sigma_t dW_t \\ d \log \Lambda_t &= \dots dt + \gamma_t dV_t \end{aligned} \right] P, P^*$$

$$\tilde{S}_t = \frac{S_t}{\Lambda_t} \Rightarrow \log \tilde{S}_t = \log S_t - \log \Lambda_t$$

$$\Rightarrow$$

$$\begin{aligned} (d \log \tilde{S}_t)^2 &= (d \log S_t)^2 + (d \log \Lambda_t)^2 - 2(d \log S_t)(d \log \Lambda_t) \\ &= \sigma_t^2 (dW_t)^2 + \gamma_t^2 (dV_t)^2 - 2\sigma_t \gamma_t \underbrace{(dW_t)(dV_t)}_{\rho_t dt} \end{aligned}$$

$$= (\sigma_t^2 + \gamma_t^2 - 2\sigma_t \gamma_t \rho_t) dt$$

$$\rho_t = \frac{1}{dt} d[W, V]_t = \text{correlation } (dW, dV)$$

(can be random)

$$\tilde{\sigma}_t^2 = \sigma_t^2 + \gamma_t^2 - 2\sigma_t \gamma_t \rho_t \quad P, P^*$$

$$\neq \sigma_t^2 \text{ ex } \gamma_t = 0 : \text{ MONEY MARKET BOND}$$

OTHER NUMERAIRE

$$d \log S_t = \dots dt + \sigma_t dW_t$$

$$d \log \Lambda_t = \dots dt + \gamma_t dV_t$$

so $\underbrace{d \log \tilde{S}_t = \dots dt + \sigma_t + \sigma_t dW_t - \gamma_t dV_t}_{\left[= \dots dt + \tilde{\sigma}_t d\tilde{W}_t \right]}$ previous page

or $d\tilde{S}_t = \underbrace{\dots \tilde{S}_t dt + \tilde{\sigma}_t \tilde{S}_t d\tilde{W}_t}_{= 0 \text{ if } P^* \text{ for numeraire } \Lambda}$

$$\tilde{\sigma}_t d\tilde{W}_t = \sigma_t dW_t - \gamma_t dV_t$$

or

$$d\tilde{W}_t = \frac{\sigma_t}{\tilde{\sigma}_t} dW_t - \frac{\gamma_t}{\tilde{\sigma}_t} dV_t$$

Yet another Brownian motion

P^* DEPENDS ON NUMERAIRE

Λ_t, B_t : NUMERAIRES, SUPPOSE P^* SAME

$$M_t = \frac{\Lambda_t}{B_t} = P^* - \text{MG}$$

$$\frac{1}{M_t} = \frac{B_t}{\Lambda_t} = P^* - \text{MG}$$

$$\underbrace{d \frac{1}{M_t}}_{d\text{MG}} = \underbrace{-\frac{1}{M_t^2} dM_t}_{d\text{MG}} + \underbrace{\frac{1}{M_t^3} d[M, M]_t}_{d\text{MG??}}$$

NEED $\underbrace{d[M, M]_t}_{dt \text{ term}} = d\text{MG}$

UNIQUENESS OF DOOB-MEYER \Rightarrow

$$[M, M]_t = 0$$

$\Rightarrow M_t$ constant if continuous

$\Rightarrow \Lambda, B$ are the same

THE RADON-NIKODYM THEOREM

Theorem. *Let P, Q be probabilities, $Q \ll P$. Then $\exists r.v. \frac{dQ}{dP}$:*

$$\text{if } E_Q|X| < \infty, \text{ then } E_Q X = E_P \left(X \frac{dQ}{dP} \right).$$

$\frac{dQ}{dP}$ is unique P - a.s.

Proof and elaborations: Billingsley, Section 32.

Uniqueness: Suppose both Y, Z satisfy conditions on $\frac{dQ}{dP}$.
Set

$$A = \{\omega : Y(\omega) > Z(\omega)\}.$$

Then

$$Q(A) = E_P Y I_A > E_P Z I_A = Q(A)$$

unless $P(A) = 0$.

The finite case: $\mathcal{F} = \sigma(\mathcal{P})$:

$$\frac{dQ}{dP}(\omega) = \frac{Q(A)}{P(A)}$$

when $\omega \in A \in \mathcal{P}$

CONTINUOUS DISTRIBUTIONS

$$P(Z \in A) = \int_A f(z) dz \quad Q(Z \in A) = \int_A g(z) dz.$$

Then:

$$\frac{dQ}{dP}(\omega) = \frac{g(Z)}{f(Z)}.$$

Proof:

$$\begin{aligned} E_P \frac{g(Z)}{f(Z)} I(Z \in A) &= \int \frac{g(z)}{f(z)} I(z \in A) f(z) dz \\ &= \int I(z \in A) g(z) dz \\ &= Q(A). \end{aligned}$$

Generally: Z_1, \dots, Z_p

$$P(A) = \int_A f(z_1, \dots, z_p) dz_1 \dots dz_p$$

$$Q(A) = \int_A g(z_1, \dots, z_p) dz_1 \dots dz_p$$

$$\frac{dQ}{dP}(\omega) = \frac{g(Z_1, \dots, Z_p)}{f(Z_1, \dots, Z_p)}.$$

NORMAL PROCESSES WITH DRIFT

$$P_\theta : \quad X_{t+1} = X_t + \theta_t + \sigma \epsilon_{t+1}$$

$$\epsilon_{t+1} \sim N(0, 1), \quad \Pi \mathcal{F}_t.$$

$$f_\theta(x_{t+1} \mid \mathcal{F}_t) = \frac{1}{\sqrt{2\pi} \sigma} \exp \left\{ -\frac{(x_{t+1} - X_t - \theta_t)^2}{2\sigma^2} \right\}$$

$$f_\theta(x_1, \dots, x_t) = (2\pi\sigma^2)^{-\frac{t}{2}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{u=0}^{t-1} (x_{u+1} - x_u - \theta_u)^2 \right\}$$

$$\frac{dP_\theta}{dP_0} = \frac{f_\theta(X_1, \dots, X_t)}{f_0(X_1, \dots, X_t)}$$

$$= \exp \left\{ -\frac{1}{2\sigma^2} \sum_{u=0}^{t-1} (X_{u+1} - X_u - \theta_u)^2 + \frac{1}{2\sigma^2} \sum_{u=0}^{t-1} (X_{u+1} - X_u)^2 \right\}$$

$$= \exp \left\{ \frac{1}{\sigma^2} \sum_{u=0}^{t-1} \theta_u (X_{u+1} - X_u) - \frac{1}{2\sigma^2} \sum_{u=0}^{t-1} \theta_u^2 \right\}$$

$$= \exp \left\{ Y_t - \frac{1}{2} [Y, Y]_t \right\}$$

where: $Y_t = \frac{1}{\sigma^2} \int_0^t \theta_u dX_u$ and $\theta_u = \theta_t$ for $t < u \leq t + 1$

GIRSANOV'S THEOREM

$$\underbrace{dX_t}_{\substack{| \\ P - BM}} = \underbrace{dW_t + \theta_t dt}_{\substack{| \\ Q - BM}}$$

$$\frac{dQ}{dP} = \exp(M_T - \frac{1}{2}[M, M]_T)$$

$$\text{where } M_t = \int_0^t \theta_u dX_u = P - \text{MG}$$

$$\frac{dP}{dQ} = \exp(-M_T + \frac{1}{2}[M, M]_T)$$

$$= \exp(-\int_0^T \theta_u [dW_u + \frac{1}{2}\theta_u du] + \frac{1}{2} \int_0^T \theta_u^2 du)$$

$$= \exp(-\int_0^T \theta_u dW_u - \frac{1}{2} \int_0^T \theta_u^2 du)$$

$$= \exp(\widetilde{M}_T - \frac{1}{2}[\widetilde{M}, \widetilde{M}]_T) \text{ where } \widetilde{M}_t = -\int_0^t \theta_u dW_u$$

$Q - \text{MG}$ as function of T

SMALL INCREMENTS, θ CONSTANT

$$P_\theta : X_{t+\Delta} = X_t + \theta\Delta + \sigma(B_{t+\Delta} - B_t)$$

θ replaced by $\theta\Delta$ σ replaced by $\sigma\sqrt{\Delta}$

$$\begin{aligned} \frac{dP_\theta}{dP_0} &= \exp \left\{ \frac{\theta\Delta}{\sigma^2\Delta} X_{N\Delta} - \frac{1}{2} \frac{\theta^2\Delta^2}{\sigma^2\Delta} N \right\} \\ &= \exp \left\{ \frac{\theta}{\sigma^2} X_T - \frac{1}{2} \frac{\theta^2}{\sigma^2} T \right\} \quad T = \Delta N \end{aligned}$$

If T fixed, $N \rightarrow \infty$, $\Delta \rightarrow 0$:

$$P_\theta : X_t = \theta t + \sigma B_t$$

BROWNIAN MOTION WITH DRIFT

Form of $\frac{dP_\theta}{dP_0}$: simplest case of Girsanov's Theorem.

APPLICATION: EVALUATION OF THE EURO-RUSSIAN OPTION

Recall: $V_0 = e^{-rT} E^*[\text{payoff}]$

$$\begin{aligned}
 V_0 &= e^{-rT} E^* f\left(\max_{0 \leq t \leq T} S_t\right) \\
 &= e^{-rT} E^* f\left(\exp\left(\max_{0 \leq t \leq T} \log S_t\right)\right) \\
 &= e^{-rT} E^* f\left(\exp\left(\max_{0 \leq t \leq T} \log S_t\right)\right) \\
 &= e^{-rT} E^* f\left(\exp\left(\log S_0 + \sigma \max_{0 \leq t \leq T} (\nu t + B_t^*)\right)\right) \\
 &\quad \sigma \nu = r - \frac{1}{2} \sigma^2 \\
 &= e^{-rT} E_Q f\left(\exp(\log S_0 + \sigma \max_{0 \leq t \leq T} X_t)\right) \frac{dP^*}{dQ}
 \end{aligned}$$

$$P^* : X_t = \nu t + B_t^*$$

$Q : X_t$ is Brownian motion.

$$\frac{dP^*}{dQ} = \exp\left\{\nu X_T - \frac{1}{2} \nu^2 T\right\}.$$

Problem reduced to one involving maxima of Brownian motion.

$$M_T = \max_{0 \leq t \leq T} X_t$$

$$V_0 = e^{-rT} E_Q f(S_0 \exp(\sigma M_T)) \exp\left(\nu X_T - \frac{1}{2} \nu^2 T\right)$$

$$f_{X,M}(a, b) = \frac{2(2b - a)}{\sqrt{2\pi T^3}} \exp\left\{-\frac{(2b - a)^2}{2T}\right\}$$

$$b \geq a, 0$$

$$V_0 = e^{-rT} \int \int f(S_0 \exp(\sigma b)) \exp\left(\nu a - \frac{1}{2} \nu^2 T\right) f_{X,M}(a, b) da db$$

ADDITIVE AND MULTIPLICATIVE MARTINGALES

 M_t IS A MARTINGALE

$$Z_t = \exp(\underbrace{M_t - \frac{1}{2}[M, M]_t}_{X_t}) = f(X_t)$$

Itô:

$$\begin{aligned} dZ_t &= f'(X_t)dX_t + \frac{1}{2}f''(X_t)d[X, X]_t \\ &= \exp(X_t) \underbrace{[dX_t + \frac{1}{2}d[X, X]_t]}_{dM_t} \\ &= Z_t dM_t = \text{MG} \end{aligned}$$

Since

$$\begin{aligned} &\underbrace{dX_t + \frac{1}{2}d[X, X]_t}_{dM_t - \frac{1}{2}d[M, M]_t} + \underbrace{\frac{1}{2}d[M, M]_t}_{\frac{1}{2}d[M, M]_t} = dM_t \end{aligned}$$

In particular:

$$EZ_T = EZ_0 = E1 = 1$$

THE DISTRIBUTION OF HITTING TIMES

$$P^* : X_t = \nu t + \sigma W_t \text{ and } Q : X_t = \sigma W_t^*$$

$$\left. \frac{dP^*}{dQ} \right|_t = \exp \left\{ \frac{\nu}{\sigma^2} X_t - \frac{1}{2} \frac{\nu^2}{\sigma^2} t \right\}.$$

Martingale property:

$$\left. \frac{dP^*}{dQ} \right|_\tau = \exp \left\{ \frac{\nu}{\sigma^2} X_\tau - \frac{1}{2} \frac{\nu^2}{\sigma^2} \tau \right\}.$$

A special case

$$\tau = \inf \{t : X_t = b\}.$$

$$\begin{aligned} P^*(\tau \leq u) &= E^* I_{\{\tau \leq u\}} \\ &= E_Q I_{\{\tau \leq u\}} \exp \left\{ \frac{\nu}{\sigma^2} X_{\tau \wedge u} - \frac{1}{2} \frac{\nu^2}{\sigma^2} (\tau \wedge u) \right\} \\ &= E_Q I_{\{\tau \leq u\}} \exp \left\{ \frac{\nu}{\sigma^2} b - \frac{1}{2} \frac{\nu^2}{\sigma^2} \tau \right\} \\ &= \int_0^u \exp \left(\frac{\nu}{\sigma^2} b - \frac{1}{2} \frac{\nu^2}{\sigma^2} t \right) f_{Q,\tau}(t) dt \end{aligned}$$

$$P^*(\tau \leq u) = \int_0^u \exp\left(\frac{\nu}{\sigma^2}b - \frac{1}{2}\frac{\nu^2}{\sigma^2}t\right) f_{Q,\tau}(t) dt$$

or:

$$f_{P^*,\tau}(t) = \exp\left(\frac{\nu}{\sigma^2}b - \frac{1}{2}\frac{\nu^2}{\sigma^2}t\right) f_{Q,\tau}(t).$$

From distribution of maximum:

$$f_{Q,\tau}(t) = \frac{|b|}{\sqrt{2\pi\sigma^2 t^3}} \exp\left\{-\frac{b^2}{2\sigma^2 t}\right\}, \quad t \geq 0$$

EXAMPLE: DOWN AND IN OPTIONS

$$\eta = \begin{cases} (S_T - K)^+ & \text{of } \min_{0 \leq t \leq T} S_t \leq K' \\ 0 & \text{otherwise} \end{cases}$$

$$X_t = \log S_t - \log S_0 \quad b = \log K' - \log S_0$$

$$\begin{aligned} \text{price} &= E^* e^{-rT} \eta \\ &= E^* E^*[e^{-rT} \eta \mid \mathcal{F}_\tau] I_{(\tau \leq T)} \\ &= E^* e^{-r\tau} E^* \left[e^{-r(T-\tau)} \eta \mid \mathcal{F}_\tau \right] I_{(\tau \leq T)} \\ &= E^* e^{-r\tau} BS(T - \tau) I_{(\tau \leq T)} \\ &= \int_0^T e^{-rt} BS(T - t) f_{P^*,\tau}(t) dt \end{aligned}$$