

STOCHASTIC INTEGRALS

$$X_t = \begin{array}{l} \text{CONTINUOUS} \\ \text{PROCESS} \end{array} \begin{cases} S_t : & \text{STOCK PRICE} \\ M_t : & \text{MG} \\ W_t : & \text{BROWNIAN MOTION} \end{cases}$$

$\theta_t =$ PORTFOLIO: # X_t HELD AT t

DISCRETE TIME: $0 = t_0 < t_1 < \dots < t_n = t$:

$$P/L_t = \sum_{i < n} \theta_{t_i} \underbrace{(X_{t_{i+1}} - X_{t_i})}_{\Delta X_{t_i}}$$

GRID BECOMES “DENSE”: $\max_i \Delta t_i \rightarrow 0$

$$P/L_t \rightarrow \int_0^t \theta_u dX_u$$

INTEGRAL DEFINED AS LIMIT OF SUMS

PROPERTIES MOSTLY FROM SUMS:

$$\underbrace{\sum_{i < n} (a\theta_{t_i} + b\eta_{t_i}) \Delta X_{t_i}}_{\downarrow} = a \underbrace{\sum_{i < n} \theta_{t_i} \Delta X_{t_i}}_{\downarrow} + b \underbrace{\sum_{i < n} \eta_{t_i} \Delta X_{t_i}}_{\downarrow}$$

$$\int_0^t (a\theta_u + b\eta_u) dX_u = a \int_0^t \theta_u dX_u + b \int_0^t \eta_u dX_u$$

⇒ LINEARITY OK

TIME VARYING INTEGRAL:

$$\int_0^t \theta_u dX_u = \text{limit of}$$

$$\sum_{t_{i+1} \leq t} \theta_{t_i} (X_{t_{i+1}} - X_{t_i})$$

Limit in probability

MARTINGALE PROPERTY:

If $X_t = M_t = \text{MG}$:

$$U_t^{(n)} = \sum_{t_{i+1} \leq t} \theta_{t_i} \Delta X_{t_i} : \quad \Delta U_{t_{i+1}}^{(n)} = \theta_{t_i} \Delta X_{t_i}$$

If on grid t_0, t_1, \dots :

$$\begin{aligned} E(\Delta U_{t_{i+1}}^{(n)} \mid \mathcal{F}_{t_i}) &= E(\theta_{t_i} \Delta X_{t_i} \mid \mathcal{F}_{t_i}) \\ &= \theta_{t_i} E(\Delta X_{t_i} \mid \mathcal{F}_{t_i}) = 0 \end{aligned}$$

$\Rightarrow U_t^{(n)}$ is \mathcal{F}_{t_i} - MG

Taking limits:

$$E(U_t \mid \mathcal{F}_s) = U_s$$

QUADRATIC VARIATION (Q. V.)

$$U_t^{(n)} = \sum_{t_{i+1} \leq t} \theta_{t_i} \Delta X_{t_i}$$

$$\text{so: } \Delta U_{t_i}^{(n)} = U_{t_{i+1}}^{(n)} - U_{t_i}^{(n)} = \theta_{t_i} \Delta X_{t_i}$$

$$(\Delta U_{t_i}^{(n)})^2 = \theta_{t_i}^2 (\Delta X_{t_i})^2$$

Aggregate:

$$\begin{aligned} \underbrace{[U^{(n)}, U^{(n)}]_t}_{\downarrow} &= \sum_{t_{i+1} \leq t} (\Delta U_{t_i}^{(n)})^2 \\ &= \sum_{t_{i+1} \leq t} \theta_{t_i}^2 (\Delta X_{t_i})^2 = \underbrace{\sum_{t_{i+1} \leq t} \theta_{t_i}^2 \Delta[X, X]_{t_i}}_{\downarrow} \\ [U, U]_t &= \int_0^t \theta_u^2 d[X, X]_u \end{aligned}$$

If $X_t = W_t = \text{B.M.} : d[X, X]_t = dt$

IT FOLLOWS THAT $[U, U]_t = \int_0^t \theta_u^2 du$

DIFFERENTIAL NOTATION

INTEGRAL:

$$\Delta U_{t_i}^{(n)} = \theta_{t_i} \Delta X_{t_i} \text{ vs. } U_t = U_0 + \sum_{t_{i+1} \leq t} \theta_{t_i} \Delta X_{t_i}$$

becomes

$$dU_t = \theta_t dX_t \text{ vs. } U_t = U_0 + \int_0^t \theta_s dX_s$$

QUADRATIC VARIATION:

$$(\Delta U_{t_i}^{(n)})^2 = \theta_{t_i}^2 (\Delta X_{t_i})^2 \text{ vs. } [U^{(n)}, U^{(n)}]_t = \sum \theta_{t_i}^2 \Delta[X, X_{t_i}]$$

” ”

$$\Delta[U^{(n)}, U^{(n)}]_{t_i} = \theta_{t_i}^2 \Delta[X, X]_{t_i}$$

becomes:

$$(dU_t)^2 = \theta_t^2 (dX_t)^2 \text{ vs. } [U, U]_t = \int_0^t \theta_u^2 d[X, X]_u$$

” ”

$$d[U, U]_t = \theta_t^2 d[X, X]_t$$

BROWNIAN MOTION:

$$(dW_t)^2 = dt \quad \text{AND} \quad d[U, U]_t = \theta_t^2 dt$$

QUADRATIC COVARIATION:

$$U, Z : \quad [U, Z]_t = \text{limit of } \sum_{t_{i+1} \leq t} \Delta U_{t_i} \Delta Z_{t_i}$$

CASE OF TWO INTEGRALS:

$$U_t = \int_0^t \theta_s dX_s, \quad Z_t = \int_0^t \eta_s dY_s$$

THEN:

$$[U, Z]_t = \int_0^t \theta_s \eta_s d[X, Y]_s$$

BECAUSE

$$\Delta U_{t_i} \Delta Z_{t_i} = \theta_{t_i} \eta_{t_i} \Delta X_{t_i} \Delta Y_{t_i}$$

or

$$\underbrace{dU_t dZ_t}_{d[U, Z]_t} = \theta_t \eta_t \underbrace{dX_t dY_t}_{d[X, Y]_t}$$

IF: $X_t = Y_t = W_t$ THE SAME B.M.:

$$d[U, Z]_t = \theta_t \eta_t dt$$

DETERMINISTIC INTEGRAND

IF θ_t IS NONRANDOM:

$$\int_0^t \theta_s dW_s = \text{limit of } \sum_{t_{i+1} \leq t} \theta_{t_i} \Delta W_{t_i}$$

$$\sum_{t_{i+1} \leq t} \theta_{t_i} \Delta W_{t_i}:$$

- LINEAR COMBINATION OF NORMAL RANDOM VARIABLES IS A NORMAL RANDOM VARIABLE

- MEAN: $E \sum_{t_{i+1} \leq t} \theta_{t_i} \Delta W_{t_i} = 0$

- VARIANCE:

$$\text{Var} \left(\sum_{t_{i+1} \leq t} \theta_{t_i} \Delta W_{t_i} \right) = \sum_{t_{i+1} \leq t} \theta_{t_i}^2 \text{Var} (\Delta W_{t_i}) = \sum_{t_{i+1} \leq t} \theta_{t_i}^2 \Delta t_i$$

IN THE LIMIT:

$$\int_0^t \theta_s dW_s:$$

- NORMAL RANDOM VARIABLE

- MEAN IS ZERO

- VARIANCE:

$$\text{Var} \left(\int_0^t \theta_s dW_s \right) = E \left[\int_0^t \theta_s dW_s, \int_0^t \theta_s dW_s \right]_t = \int_0^t \theta_s^2 ds$$

ITÔ'S FORMULA

X_t : CONTINUOUS PROCESS (SOME RESTRICTIONS):
 ξ : TWICE CONTINUOUSLY DIFFERENTIABLE

$$\xi(X_t) = \xi(X_0) + \int_0^t \xi'(X_u) dX_u + \frac{1}{2} \int_0^t \xi''(X_u) d[X, X]_t$$

DIFFERENTIAL NOTATION:

$$d\xi(X_t) = \xi'(X_t) dX_t + \frac{1}{2} \xi''(X_t) d[X, X]_t$$

—

EX: $X_t = W_t =$ BROWNIAN MOTION:

$$d\xi(W_t) = \xi'(W_t) dW_t + \frac{1}{2} \xi''(W_t) dt$$

—

EX: $dX_t = \nu_t dt + \sigma_t dW_t$ ITÔ PROCESS

or: $X_t = X_0 + \int_0^t \nu_s ds + \int_0^t \sigma_s dW_s$

First question: ¿What is $d[X, X]_t$?

FIRST: CASE OF EXPLICIT INTEGRATION:

$W_t = \text{B.M.}$ ¿WHAT IS $\int_0^t W_s dW_s$?

$$U_t = W_t^2 = \zeta(W_t), \zeta(x) = x^2$$

$$\begin{aligned} dU_t &= \zeta'(W_t)dW_t + \frac{1}{2}\zeta''(W_t)dt \\ &= 2W_t dW_t + dt \end{aligned}$$

$$\text{so: } W_t dW_t = \frac{1}{2}dU_t - \frac{1}{2}dt$$

$$\begin{aligned} >: \int_0^t W_s dW_s &= \frac{1}{2}(U_t - U_0) - \frac{1}{2} \int_0^t ds \\ &= \frac{1}{2}W_t^2 - \frac{1}{2}t \end{aligned}$$

DIFFERENT FROM ORDINARY INTEGRAL:

If $X_t = g(t)$ g' exists, continuous, $g(0) = 0$

$$\begin{aligned} \int_0^t X_s dX_s &= \int_0^t g(s)g'(s)ds \\ &= \frac{1}{2}g(t)^2 \\ &= \frac{1}{2}X_t^2 \end{aligned}$$

ITÔ PROCESS:

$$X_t = X_0 + \underbrace{\int_0^t \nu_s ds}_{Z_t} + \underbrace{\int_0^t \sigma_s dW_s}_{U_t}$$

Grid:

$$\Delta X_{t_i} = \Delta Z_{t_i} + \Delta U_{t_i}$$

SO:

$$(\Delta X_{t_i})^2 = (\Delta Z_{t_i})^2 + (\Delta U_{t_i})^2 + 2\Delta Z_{t_i} \Delta U_{t_i}$$

$$\sum (\Delta X_{t_i})^2 = \sum (\Delta Z_{t_i})^2 + \sum (\Delta U_{t_i})^2 + \sum 2\Delta Z_{t_i} \Delta U_{t_i}$$

$$\begin{aligned} |\Delta Z_{t_i}| &= \left| \int_{t_i}^{t_{i+1}} \nu_s ds \right| \leq \int_{t_i}^{t_{i+1}} |\nu_s| ds \\ &\leq \sup_s |\nu_s| (t_{i+1} - t_i) = \sup_s |\nu_s| \Delta t_i \end{aligned}$$

$$\begin{aligned} \sum (\Delta Z_{t_i})^2 &\leq (\sup_s |\nu_s|)^2 \sum_i (\Delta t_i)^2 \\ &\leq (\sup_s |\nu_s|)^2 \underbrace{\sup_i \Delta t_i}_{\rightarrow 0} \underbrace{\sum_i \Delta t_i}_{\simeq t} \rightarrow 0 \end{aligned}$$

$$\subset: [Z, Z]_t = 0 \quad \text{ALSO: } [Z, U]_t = 0$$

$$\text{ONLY: } [U, U]_t = \int_0^t \sigma_s^2 ds$$

ITÔ PROCESS

$$X_t = X_0 + \underbrace{\int_0^t \nu_s ds}_{Z_t} + \underbrace{\int_0^t \sigma_s dW_s}_{U_t}$$

$$d[Z, Z]_t = 0 \quad d[Z, U]_t = 0 \quad d[U, U]_t = \sigma_t^2 dt$$

$$(dZ_t)^2 \quad dZ_t dU_t$$

USING DIFFERENTIALS:

ANY dt -TERM HAS ZERO Q.V.:

$$(dZ_t)^2 = \nu_t^2 (dt)^2 = 0 \text{ ETC}$$

—

COMBINING TERMS:

$$\begin{aligned} (dX_t)^2 &= (dZ_t + dU_t)^2 \\ &= (dZ_t)^2 + 2dZ_t dU_t + (dU_t)^2 \\ &= (dU_t)^2 = \sigma_t^2 dt \end{aligned}$$

RIGOROUS:

$$(\Delta X_{t_i})^2 = \Delta Z_{t_i}^2 + 2\Delta Z_{t_i} \Delta U_{t_i} + (\Delta U_{t_i})^2$$

SUM OVER t_i , TAKE LIMITS, GET SAME RESULT

INTEGRALS WITH RESPECT TO AN ITÔ PROCESS

$$X_t = X_0 + \int_0^t \nu_s ds + \int_0^t \sigma_s dW_s$$

CAN SHOW THAT:

$$\int_0^t \theta_s dX_s = \int_0^t \theta_s \nu_s ds + \int_0^t \theta_s \sigma_s dW_s$$

A NEW ITÔ PROCESS

BACK TO ITÔ'S FORMULA:

$$d\xi(X_t) = \xi'(X_t)dX_t + \frac{1}{2}\xi''(X_t)d[X, X]_t \quad (*)$$

ITÔ PROCESS:

$$dX_t = \nu_t dt + \sigma_t dW_t$$

SO

$$d[X, X]_t = \sigma_t^2 dt$$

PLUG IN:

$$\begin{aligned} d\xi(X_t) &= \xi'(X_t)(\nu_t dt + \sigma_t dW_t) \\ &\quad + \frac{1}{2}\xi''(X_t)\sigma_t^2 dt \\ &= (\xi'(X_t)\nu_t + \frac{1}{2}\xi''(X_t)\sigma_t^2)dt \\ &\quad + \xi'(X_t)\sigma_t dW_t \end{aligned}$$

EASIER TO REMEMBER (*). . .

“PROOF” OF ITÔ’S FORMULA:

$$U_t = \xi(X_t):$$

$$\begin{aligned} \Delta U_{t_i} &= \xi(X_{t_{i+1}}) - \xi(X_{t_i}) \\ &= \xi(X_{t_i} + \Delta X_{t_i}) - \xi(X_{t_i}) \\ &= \zeta'(X_{t_i})\Delta X_{t_i} + \frac{1}{2}\zeta''(X_{t_i})\Delta X_{t_i}^2 \\ &\quad + \frac{1}{3!}\zeta'''(X_{t_i})\Delta X_{t_i}^3 + \dots \end{aligned}$$

sum up:

$$\begin{aligned} U_t - U_0 &= \underbrace{\sum \zeta'(X_{t_i})\Delta X_{t_i}}_{\downarrow} + \underbrace{\frac{1}{2}\sum \zeta''(X_{t_i})\Delta X_{t_i}^2}_{\downarrow} \\ &= \int_0^t \zeta'(X_s)dX_s + \frac{1}{2}\int_0^t \zeta''(X_s)d[X, X]_s \end{aligned}$$

OTHER “PROOF”:

$$\begin{aligned} dU_t &= \zeta(X_t + dX_t) - \zeta(X_t) \\ &= \zeta'(X_t)dX_t + \frac{1}{2}\zeta''(X_t)\underbrace{(dX_t)^2}_{d[X, X]_t} + \dots \end{aligned}$$

MULTIVARIATE FORMULA

$$\begin{aligned}U_t &= \zeta(X_t, Y_t) \\dU_t &= \zeta'_x(X_t, Y_t)dX_t + \zeta'_y(X_t, Y_t)dY_t \\&\quad + \frac{1}{2} \left\{ \zeta''_{xx}(X_t, Y_t)d[X, X]_t \right. \\&\quad \quad + \zeta''_{yy}(X_t, Y_t)d[Y, Y]_t \\&\quad \quad \left. + 2\zeta''_{xy}(X_t, Y_t)d[X, Y]_t \right\}\end{aligned}$$

etc.

EXAMPLE: GEOMETRIC BROWNIAN MOTION

$$S_t = S_0 \exp\left\{\int_0^t \sigma_s dW_s + \int_0^t \left(r_s - \frac{1}{2}\sigma_s^2\right) ds\right\}$$

SET

- $X_t = \int_0^t \sigma_s dW_s + \int_0^t \left(r_s - \frac{1}{2}\sigma_s^2\right) ds$
- $S_t = f(X_t)$ ($f(x) = S_0 \exp\{x\}$)

USE ITÔ'S FORMULA

$$dS_t = f'(X_t)dX_t + \frac{1}{2}f''(X_t)d[X, X]_t$$

$f'(x) = f''(x) = f(x)$ AND $d[X, X]_t = \sigma_t^2 dt$, SO:

$$\begin{aligned} dS_t &= f(X_t)dX_t + \frac{1}{2}f(X_t)\sigma_t^2 dt \\ &= S_t dX_t + \frac{1}{2}S_t \sigma_t^2 dt \\ &= S_t \left(dX_t + \frac{1}{2}\sigma_t^2 dt \right) \\ &= S_t (\sigma_t dW_t + r_t dt) \\ &= S_t \sigma_t dW_t + S_t r_t dt \end{aligned}$$

DIFFERENTIAL REPRESENTATION OF S_t

VASICEK MODEL

$$dR_t = (\alpha - \beta R_t)dt + \sigma dW_t$$

- STEP 1: SET $U_t = R_t - \frac{\alpha}{\beta}$

EQUATION BECOMES:

$$dU_t = -\beta U_t dt + \sigma dW_t$$

- STEP 2: NOTE THAT (FROM ITO'S FORMULA)

$$\begin{aligned} d(\exp\{\beta t\}U_t) &= \exp\{\beta t\}dU_t + U_t d\exp\{\beta t\} \\ &= \exp\{\beta t\}dU_t + U_t\beta \exp\{\beta t\}dt \\ &= \exp\{\beta t\} (dU_t + U_t\beta dt) \\ &= \exp\{\beta t\}\sigma dW_t \end{aligned}$$

SO

$$\exp\{\beta t\}U_t = U_0 + \int_0^t \exp\{\beta s\}\sigma dW_s$$

OR

$$U_t = \exp\{-\beta t\}U_0 + \int_0^t \exp\{\beta(s-t)\}\sigma dW_s$$

IN OTHER WORDS: U_t IS NORMAL

- MEAN IS $\exp\{-\beta t\}U_0$
- VARIANCE IS

$$\begin{aligned} & \int_0^t (\exp\{\beta(s-t)\}\sigma)^2 ds \\ &= \int_0^t \exp\{2\beta(s-t)\}\sigma^2 ds \\ &= \left[\frac{1}{2\beta} \exp\{2\beta(s-t)\}\sigma^2 \right]_{s=0}^{s=t} \\ &= \frac{1}{2\beta} (1 - \exp\{-2\beta t\}) \sigma^2 \end{aligned}$$

DEDUCE FOR $R_t = U_t + \frac{\alpha}{\beta}$ THAT

- R_t IS NORMAL
- $E(R_t) = \exp\{-\beta t\}U_0 + \frac{\alpha}{\beta}$
- $\text{Var}(R_t) = \text{Var}(U_t)$

LEVY'S THEOREM

IF M_t IS A CONTINUOUS (LOCAL) MARTINGALE,
 $M_0 = 0$, $[M, M]_t = t$ FOR ALL t , THEN M_t IS A
 CONTINUOUS BROWNIAN MOTION

PROOF: SET $f(x) = \exp\{hx\}$

ITO:

$$\begin{aligned} df(M_t) &= f'(M_t)dM_t + \frac{1}{2}f''(M_t)d[M, M]_t \\ &= f'(M_t)dM_t + \frac{1}{2}f''(M_t)dt. \end{aligned}$$

SINCE dM_t TERM IS MG, AND $f''(x) = h^2 f(x)$:

$$\begin{aligned} E(f(M_t)|_s) &= f(M_s) + \frac{1}{2}h^2 E\left(\int_s^t f(M_u)du|_s\right) \\ &= f(M_s) + \frac{1}{2}h^2 \int_s^t E(f(M_u)|_s)du \end{aligned}$$

Set $g(t) = E(\exp\{h(M_t - M_s)\}|_s)$:

$$g(t) = 1 + \frac{1}{2}h^2 \int_s^t g(u)du$$

SOLUTION:

$$g(t) = \exp\left\{\frac{1}{2}h^2(t - s)\right\}$$

IN OTHER WORDS:

$$E(\exp\{h(M_t - M_s)\} | \mathcal{F}_s) = \exp\left\{\frac{1}{2}h^2(t - s)\right\}$$

CHARACTERISTIC FUNCTION ARGUMENT GIVES:

- $M_t - M_s$ IS INDEPENDENT OF \mathcal{F}_s
- $M_t - M_s$ IS $N(0, t - s)$

ITÔ PROCESSES

$$X_t = X_0 + \underbrace{\int_0^t a_s ds}_{dt \text{ term}} + \underbrace{\int_0^t b_s dW_s}_{dW_t \text{ term}}$$

Decomposition unique:

- dW_t term is martingale
- dt term is drift

(Doob-Meyer decomposition)

UNDER RISK NEUTRAL MEASURE P^*

Discounted securities only have dW term:

$$d\tilde{S}_t = \underbrace{\mu_t \tilde{S}_t dt}_{= 0} + \sigma_t \tilde{S}_t dW_t$$

UNDISCOUNTED SECURITIES UNDER P^* :

$$d\tilde{S}_t = \sigma_t \tilde{S}_t dW_t$$

1) Numeraire = $B_t = \exp(\int_0^t r_u du)$

properties: $dB_t = r_t B_t dt$, $[B, B]_t = 0$, $[B, \tilde{S}]_t = 0$

$$\begin{aligned} S_t &= \tilde{S}_t B_t \Rightarrow \\ dS_t &= B_t d\tilde{S}_t + \tilde{S}_t dB_t \\ &= \underbrace{B_t \sigma_t \tilde{S}_t dW_t}_{\swarrow \searrow} + \underbrace{\tilde{S}_t B_t r_t dt}_{\nearrow} \\ &= \sigma_t S_t dW_t + r_t S_t dt \end{aligned}$$

2) Other numeraire:

$\Lambda_t \neq B_t$, Λ_t has $dW_t^{(2)}$ term

$$d\langle W, W^{(2)} \rangle_t = \rho_t dt$$

$d\tilde{S}_t \Lambda_t =$ full use of Itô's formula

Not same $P^*!!!$

UNDISCOUNTED SECURITIES UNDER P^* :

$$dS_t = r_t S_t dt + \sigma_t S_t dW_t$$

LOG SCALE: ITO'S FORMULA

$$\begin{aligned} d \log(S_t) &= \frac{1}{S_t} dS_t + \frac{1}{2} \left(-\frac{1}{S_t^2} \right) d[S, S]_t \\ &= \frac{1}{S_t} (r_t S_t dt + \sigma_t S_t dW_t) + \frac{1}{2} \left(-\frac{1}{S_t^2} \right) \sigma_t^2 S_t dt \\ &= \left(r_t - \frac{1}{2} \sigma_t^2 \right) dt + \sigma_t dW_t \end{aligned}$$

OPTIONS PRICES: PDE'S

PAYOFF: $f(S_T)$

DISCOUNTED PAYOFF: $e^{-rT} f(S_T) = \tilde{f}(\tilde{S}_T)$

$$\tilde{S}_T = e^{-rT} S_T \quad \tilde{f}(\tilde{s}) = e^{-rT} f(e^{rT} \tilde{s})$$

CALL: $f(s) = (s - K)^+$ $\tilde{f}(\tilde{s}) = (\tilde{s} - e^{-rT} K)^+$

CANDIDATE PRICE: DISCOUNTED:

$$\tilde{C}(\tilde{S}_t, t) \text{ satisfies: } \tilde{C}(\tilde{S}, T) = \tilde{f}(\tilde{S})$$

AND (UNDER P^*):

$$d\tilde{C}(\tilde{S}_t, t)$$

$$\text{Hedge } \left\{ \begin{array}{l} = \tilde{C}'_s(\tilde{S}_t, t) d\tilde{S}_t \\ \end{array} \right\} \begin{array}{l} \text{MG term} \\ d\tilde{S} = \sigma_t \tilde{S}_t dW_t \end{array}$$

$$\text{BS PDE } \left\{ \begin{array}{l} + \tilde{C}'_t(\tilde{S}_t, t) dt \\ + \frac{1}{2} \tilde{C}''_{ss}(\tilde{S}_t, t) \underbrace{d[\tilde{S}, \tilde{S}]_t}_{\sigma_t^2 \tilde{S}_t^2 dt} \end{array} \right\} dt - \text{terms} = 0$$

2 APPROACHES

THE BS PDE:

$$(*) \quad \begin{cases} \tilde{C}'_t(\tilde{s}, t) + \frac{1}{2} \tilde{C}''_{ss}(\tilde{s}, t) \sigma^2 \tilde{s}^2 = 0 \\ \tilde{C}(\tilde{s}, T) = \tilde{f}(\tilde{s}) \end{cases}$$

THE MARTINGALE APPROACH:

Set

$$\tilde{C}(\tilde{s}, t) = E^*[\tilde{f}(\tilde{S}_T) \mid \tilde{S}_t = \tilde{s}]$$

Markov: $\tilde{C}(\tilde{S}_t, t) = E^*[\tilde{f}(\tilde{S}_T) \mid \mathcal{F}_t] = \text{price under } P^*$

This \tilde{C} either

- 1) Market is complete:
 \tilde{C} automatically satisfies (*)
- 2) Otherwise: check if \tilde{C} satisfies (*):
 yes: solution OK
 no: try something else

REVERSAL OF DISCOUNTING

Numeraire: $B_t = \exp\{rt\}$

$$\begin{aligned}C(S_t, t) &= B_t \tilde{C}(\tilde{S}_t, t) \\ &= B_t \tilde{C}\left(\frac{S_t}{B_t}, t\right) \\ &= e^{rt} \tilde{C}(e^{-rt} S_t, t)\end{aligned}$$

$$\begin{aligned}\text{Hence: } C(s, t) &= e^{rt} \tilde{C}(e^{-rt} s, t) \\ &= e^{rt} E^*[\tilde{f}(\tilde{S}_T) \mid \tilde{S}_t = e^{-rt} s] \\ &= e^{r(T-t)} E^*[f(S_T) \mid S_t = s]\end{aligned}$$

since

$$\tilde{f}(\tilde{s}) = e^{-rT} f(e^{rT} \tilde{s})$$

COMPUTATION OF EXPECTED VALUES

$$\log \tilde{S}_T = \log \tilde{S}_t + \underbrace{\nu(T-t)}_{\nu = -\frac{\sigma^2}{2}} + \sigma \underbrace{(W_T - W_t)}_{\substack{\sqrt{T-t} Z \\ Z \sim N(0,1)}}$$

and so : $\tilde{S}_T = \tilde{S}_t \exp(\nu(T-t) + \sigma\sqrt{T-t} Z)$

$$\begin{aligned} \tilde{C}(s, t) &= E[\tilde{f}(\tilde{S}_T) \mid \tilde{S}_t = \tilde{s}] \\ &= E[\tilde{f}(\tilde{s} \exp(\nu(T-t) + \sigma\sqrt{T-t} Z))] \\ &= \int_{-\infty}^{+\infty} \tilde{f}(\tilde{s} \exp(\nu(T-t) + \sigma\sqrt{T-t} z)) \phi(z) dz \\ &\quad \text{where } \phi(z) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}z^2\right) \end{aligned}$$

For non-discounted S_t : $\nu = r - \frac{1}{2}\sigma^2$

OPTION PRICES: GENERAL SCHEME

SELF FINANCING STRATEGIES:

$$\eta = C_T$$

$$dC_t = \theta_t^{(0)} dB_t + \sum_{i=1}^K \theta_t^{(i)} dS_t^{(i)}$$

$$C_t = \theta_t^{(0)} B_t + \sum_{i=1}^K \theta_t^{(i)} S_t^{(i)}$$

Same as (by numeraire invariance):

$$\left. \begin{aligned} \tilde{\eta} &= \frac{1}{B_T} \eta & : & \quad \tilde{\eta} = \tilde{C}_T \\ d\tilde{C}_t &= \sum_{i=1}^K \theta_t^{(i)} d\tilde{S}_t^{(i)} \end{aligned} \right] \begin{array}{l} \text{Key} \\ \text{requirement} \end{array}$$

$$\left. \tilde{C}_t = \theta_t^{(0)} + \sum_{i=1}^K \theta_t^{(i)} \tilde{S}_t^{(i)} \right] \begin{array}{l} \text{Defines} \\ \theta_t^{(0)} \end{array}$$

PROOF: ITÔ'S FORMULA

ON THE DISCOUNTED SCALE

$$\left. \begin{aligned} \tilde{\eta} &= \tilde{C}_T \\ d\tilde{C}_t &= \sum_{i=1}^K \theta_t^{(i)} d\tilde{S}_t^{(i)} \end{aligned} \right]$$

UNDER P^* : SAME AS

$$\tilde{\eta} = c + \sum_{i=1}^K \int_0^T \theta_t^{(i)} d\tilde{S}_t^{(i)} \quad (*)$$

BY TAKING

$$\tilde{C}_t = E^*(\tilde{\eta} \mid \mathcal{F}_t)$$

If $B_0 = 1$: $c = \tilde{C}_0 = \text{PRICE AT } 0$

(*): “MARTINGALE REPRESENTATION THEOREM”

WHEN DOES THE REPRESENTATION THEOREM HOLD?

THEOREM: *IF* $W^{(1)}, \dots, W^{(K)}$ *INDEPENDENT B.M.'S*

$$\mathcal{F}_t = \mathcal{F}_t^{W^{(1)}, \dots, W^{(K)}}$$

$$\tilde{\eta} @ \mathcal{F}_T, \quad E^* |\tilde{\eta}| < \infty :$$

$$\tilde{\eta} = c + \sum_{i=1}^P \int_0^T f_t dW_t^{(i)}$$

Brownian motions are like binomial trees

FROM BROWNIAN MOTION TO STOCK PRICE

$$d\tilde{S}_t = \sigma \tilde{S}_t dW_t$$

or:

$$\log \tilde{S}_t = \log \tilde{S}_0 - \frac{1}{2}\sigma^2 t + \sigma \tilde{W}_t$$

$$\tilde{\eta} @ \mathcal{F}_T^{\tilde{S}} \Leftrightarrow \tilde{\eta} @ \mathcal{F}_T^W$$

$$\begin{aligned} \text{Get: } \tilde{\eta} &= c + \int_0^T f_t dW_t \\ &= c + \int_0^T \underbrace{\frac{f_t}{\sigma \tilde{S}_t}}_{\theta_t} d\tilde{S}_t \end{aligned}$$

More complex if σ_t or r_t random...